# Comparing Greenbook and Reduced Form Forecasts using a Large Realtime Dataset[*]

## Jon Faust[†]       Jonathan H. Wright[‡]

[†]Department of Economics, Johns Hopkins University, Baltimore MD 21218 and Division of International Finance, Federal Reserve Board, Washington DC 20551, phone: (202) 452-2328, e-mail: faustj@jhu.edu

[‡]Division of Monetary Affairs, Federal Reserve Board, Washington DC 20551, phone: (202) 452-3605, e-mail: jonathan.h.wright@frb.gov.

Abstract: Many recent papers have found that atheoretical forecasting methods using many predictors give better predictions for key macroeconomic variables than various small-model methods. The practical relevance of these results is open to question, however, because these papers generally use *ex-post* revised data not available to forecasters and because no comparison is made to best actual practice. We provide some evidence on both of these points using a new large dataset of vintage data synchronized with the Fed's Greenbook forecast. This dataset consists of a large number of variables, as observed at the time of each Greenbook forecast since 1979. Thus, we can compare real-time large dataset predictions to both simple univariate methods and to the Greenbook forecast. For inflation we find that univariate methods are dominated by the best atheoretical large dataset methods and that these, in turn, are dominated by Greenbook. For GDP growth, in contrast, we find that once one takes account of Greenbook's advantage in evaluating the current state of the economy, neither large dataset methods nor the Greenbook process offers much advantage over a univariate autoregressive forecast.

## 1. Introduction

In recent years, researchers have investigated many different ways of forecasting an economic time series using a large number of predictors—say, 40 or more (e.g., Stock and Watson, 1999, 2002, 2003, 2005; Bernanke, Boivin and Eliasz, 2005; Giannone, Reichlin and Sala, 2004; Forni, Hallin, Lippi and Reichlin, 2005). When the number of predictors is large relative to the available sample size, one must constrain the estimation of the forecasting model in some way in order to avoid the perils of overfitting, and the many methods differ mainly in how they do so. The methods have one aspect in common, however: they are atheoretical in the sense that restrictions from economic theory are eschewed.[1]

A consistent result in this work is that these large-dataset methods outperform various naive and semi-sophisticated benchmarks including random walk forecasts, simple univariate time series models, and, in some cases, simple models motivated by economic theory.

While these pioneering results are tantalizing, their importance for practical forecasting is difficult to assess for two related reasons. First, arguably none of the benchmarks used in this research is used for practical forecasting. While the simple benchmarks are probably the right starting point for assessing new methods, ultimately we care whether the new methods outperform standard practice or best practice in forecasting. Second, comparison to real-world forecasting methods is complicated by the issue of data revisions. Many macro time series are heavily revised through time. Real-world forecasts use the noisy early vintages available in real-time; large-dataset forecasting research has been conducted almost exclusively with a single vintage of revised data.

Bernanke and Boivin (2003) provide a notable exception. They use a real-time dataset to assess various large dataset forecasting methods and they include a comparison to the Fed's Greenbook forecast. Two results are especially interesting and provocative.

---

[1] More specifically, economics only guides the selection of variables to include in the exercise.

First, factor models generally predict less well in their 78-variable real-time dataset than using the 215 series of fully-revised data used by Stock and Watson (2002). This appears to owe mainly to the larger size of the Stock and Watson dataset, because the factor models also predict less well in the revised version of the 78-variable dataset. Second, the Greenbook seems to outperform all the other methods considered—regardless of the dataset employed or whether revised data are used.

In this paper, we take up some of the questions raised by Bernanke and Boivin (2003), using a unique set of vintage data associated with the Federal Reserve's Greenbook forecast, which is prepared for each FOMC meeting.[2] The dataset has a snapshot of a large number of macroeconomic time series as they existed at the time of 145 Greenbooks between March 1980 and December 2000. These data allow us to create pseudo-real-time large dataset forecasts using information sets that are precisely synchronized with the Greenbook.

First, we compare various large dataset forecasts of output growth and inflation to simple benchmarks in real time, to see if the conclusions are sensitive to data-vintage issues. Following Bernanke and Boivin (2003), we study factor model approaches. We also study two model averaging methods in which the overall forecast is a weighted average of forecasts from a large number of simple bivariate models. We find that the model averaging methods are both more robust and generally more accurate than the factor model methods—the factor models sometimes perform very badly.

Second, we compare the large dataset forecasts to the Greenbook forecasts, exploiting the fact that the information sets are perfectly synchronized. This comparison is particularly interesting because much earlier work suggests that the Greenbook forecast

---

[2]Other authors who have used these electronic archives in this way include Tetlow and Ironside (2004), who evaluated alternative policy rules with real-time vintages of the data and of the FRB/US model, and Edge, Kiley and Laforte (2006), who used data on 11 series from these databases since 1996 to compare Greenbook forecasts with real-time forecasts based on Bayesian VARs and dynamic stochastic general equilibrium (DSGE) models.

has generally been at or near the frontier of best performance in forecasting. We are not aware of documentation of a real-time forecast that consistently outperformed Greenbook over this period.[3] We do not mean to imply that Greenbook is known to be optimal in some sense, only that it is near the frontier of best actual practice, and thereby an interesting benchmark. Further, the Greenbook is a subjective forecast based on an immense range of information processed through a subjective filter that is guided by economics.[4] Thus, we can view the large-dataset-forecast versus Greenbook comparison as a test of the atheoretical use of a large dataset versus sophisticated use of an immense dataset.

Using this unique dataset, we can considerably sharpen the comparison of Greenbook to other methods over what has been done heretofore. In particular, we can evaluate whether the well-documented Greenbook advantages stem mainly from superior estimates of the state of the economy at the time of the forecast. When the Greenbook forecast for GDP is made in quarter $t$, very little data for quarter $t$ may be available, and some data for quarter $t-1$ will not be available. Sims (2002) has suggested that the good properties of Greenbook might flow from the great effort the Fed makes to evaluate the current or recent past state of the economy at the time of the forecast. For example, by mirroring key elements of the data construction machinery of the Bureau of Economic Analysis, the Fed staff can form a relatively precise estimate of what BEA will announce for the previous quarter's GDP even before it is announced. Further, the staff can adjust the estimate of the current state for certain large transitory events such as dock strikes, hurricanes, etc. Given the persistence in economic measures, a better estimate of persistent and transitory components of the current state may translate into forecasting advantages over moderate horizons.

To put is most starkly, one might conjecture that the Fed's forecasting advantage

---

[3]Romer and Romer (2000) found that Greenbook outperformed private sector surveys and Sims (2002) found generally favorable results, especially for Greenbook inflation forecasts. Bernanke and Boivin (2003) reach the same conclusion.

[4]For a discussion of the Greenbook forecasting process, see Reifschneider, Stockton and Wilcox (1997).

stems purely from measuring the current state, with little or no advantage over atheoretical methods in projecting what the current state implies for the future. We assess this conjecture in a simple way. Consider the vintage data available in quarter $t$, so that the released values of many data series end in quarter $t-1$ or $t-2$. We have a Fed forecast of each variable, so we can append the forecast to the actual data, and bring the data all up to quarter $t-1$, or $t$, or $t+j$. We can then use this updated dataset to form the various time series forecasts. We call the point we update the data to "the jumping-off point" for the time series methods and consider jumping-off points from quarter $t-1$ through $t+3$. If the Fed's advantage mainly flows from assessing time $t-1$ or $t$, the atheoretical methods jumping off from time $t$ might compare favorably to Greenbook.

Given the generally favorable existing results on Greenbook, our initial suspicion was that for both inflation and output growth, Greenbook would look good at very short horizons, but that the time series methods jumping off from some point a quarter or two in the future would compete favorably. This suspicion turns out to be wrong for both inflation and output, but, somewhat surprisingly, in different directions. For inflation, Greenbook generally outperforms other methods (often substantially) for all jumping off points out to $t+3$: even when the time series models know what the Fed is thinking about inflation three quarters into the future, the atheoretical methods are dominated by Greenbook in forecasting further into the future. In sharp contrast, for output growth, once we give the time series methods quarter $t$, the current state, the Greenbook advantage disappears. Perhaps more surprisingly, none of the time series methods clearly outperforms an AR(4) predicting in output growth, using jumping-off points of quarter $t$ or later.

Overall, then, for inflation and for all jumping-off points, Greenbook dominates the best large model methods, which dominate the small model methods. For output, for all jumping-off points except $t-1$, the best atheoretical methods and Greenbook perform comparably and no method clearly dominates the naive benchmark of a univariate

4

autoregression. Knowledge of interest rates, inflation, unemployment, etc., do not allow one to improve on the forecast that growth will revert at the historical rate back to the current estimate of the mean.

The plan for the remainder of this paper is as follows. Sections 2 and 3 describe our real-time dataset and methods, respectively. Section 4 contains the main results; section 5 addresses a number of additional topics. Section 6 concludes.

## 2. The Real-Time Dataset

Before each FOMC meeting, the Federal Reserve staff prepares a briefing document called the Greenbook, which contains a staff forecast of the macroeconomy. While the Greenbook forecast is subjective, a large-scale econometric model has long been one tool used in the Greenbook process. Since 1996, the FRB/US model has been used. Before that, a model called variously the FMP (Federal Reserve-MIT-Penn) or MPS (MIT-Penn-SSRC) was employed. Since 1979, the model database has been archived electronically on the date that the Greenbook is published, a few days before the FOMC meeting. These archived databases contain data on many variables of which some (but not all) are forecasted in the Greenbook. The databases contain the historical sample of each variable in the model as observed at that time, and forecasts. While the forecasts are generated from the econometric model, an add-factor is then included to force the forecasts to match those in the Greenbook, for those variables that are forecasted in the Greenbook. Thus these archives comprise a large vintage dataset that is perfectly synchronized with the Greenbook forecast itself.

We extracted 145 vintages of data from these electronic archives covering Greenbooks from March 1980 to December 2000. The data used in this paper stop in 2000 because it is Federal Reserve policy that the Greenbook forecast cannot be released until 5 years after the forecast date. A few vintages are lost or never existed, including those during early 1996 during the transition from the MPS model to the FRB/US model. The

5

maximum forecast horizon varies from one Greenbook to the next, but we considered only Greenbooks for which the forecast horizon goes out to at least 5 quarters. Table 1 gives the publication dates for our 145 vintages of data.

Ideally, we might like to have the same forecast variables in each vintage. Unfortunately, the variables available in any given vintage vary a good bit. There is a large break in the list of available variables at the time of the major model revision in 1996. Finally, the amount of historical data for some series varies a bit across vintages. Further, the nature of the model databases presents an additional variable selection problem. Many variables in the databases are constructed from other variables based on identities, other transforms of included variables, and slightly different versions of included variables (e.g., foreign sector variables on both the NIPA and BOP basis).

We extracted 112 macro time series, listed in Table 2, from these databases. We omitted variables that are constructed from other variables and different definitions of the same variable. From each vintage, we keep any of these 112 that have historical data back to 1960Q2. We augment each vintage with the dividend-exclusive returns on the S&P500 stock index.[5]

In total, the number of series in each vintage ranges from 44 to 78, with an average of 64. This is a large number of predictors, although it falls short of the very large datasets of more than 100 variables that are used in some papers. Evidence on the marginal gain in predictive power from increasing the number of variables from (say) about 50 to over 100 is mixed. Boivin and Ng (2005) find little marginal gain in predictive power, but Bernanke and Boivin (2003) find a notable gain in predictive power from using a very large dataset.

In the end, we have 145 vintages containing a varying list of variables, each of which has history back to 1960Q2 and for which we have 5 quarter forecasts. The vast majority

---

[5] To each vintage we added only observations that would have been available at the Greenbook publication date. These data are of course not subject to revision

of our predictors are available out to the Greenbook forecast horizon. As is usual, we transformed many of the series so that the transformed series is arguably stationary. Some series are kept in levels (no transformation), others are in differences and others are in log differences. Table 2 notes the transformation that was applied to each series.

The three series that we forecast in this paper are the quarterly inflation rate as measured by the CPI, the quarterly inflation rate as measured by the GNP/GDP deflator and the quarterly real growth rate (GNP/GDP). As is standard, we use GNP during and before 1991, and GDP subsequently. The inflation rate and the growth rate are computed as $400 \log(x_t/x_{t-1})$ where $x$ is the price or output series.

Our goal is to assess different methods for forecasting these inflation and growth variables in real-time. But an important issue that arises is that these variables are in turn constantly being revised and so it is not clear what to treat as the realized data. For the national income and product accounts, the source of our forecast variables, the first data release (known as the *advance* release) comes out about one month after the end of the quarter to which the data refer. The data are then revised in the *preliminary* and *final* releases, which incorporate more source data, and are released about two and three months after the end of the quarter to which the data refer, respectively. The data then continue to get revised indefinitely, through annual and then benchmark revisions, with the latter incorporating conceptual and definitional changes. Since our objective is the comparison of Greenbook and statistical forecasts, and since Federal Reserve staff attempt to forecast variables as they are defined at the time that the forecast is made, we measure actual realized inflation and growth by the data as recorded in the real-time dataset of the Federal Reserve Bank of Philadelphia two quarters after the quarter to which the data refer. This will typically correspond to the data as recorded in the second revision of the national income and product accounts.

## 3. Methods

*3.1 The forecasting models*

We construct forecasts using 10 time series methods from the literature. Call the variable to be forecast $y_t$ (a measure of inflation or output growth) and a collection of potential predictors $\{x_{it}\}_{i=1}^n$. At time $T$ we observe data up to some point in time $T - \ell$. We will consider forecasts for $y_{T+h}$, $h = 0, \ldots, 5$.

We consider the following time series forecasts:

1. A random walk model (RW). This takes $y_T$ as the forecast for $y_{T+h}$. This is close to, but not quite the same as, the forecast for inflation considered by Atkeson and Ohanian (2001), who take $\frac{1}{4}\Sigma_{j=0}^3 y_{T-j}$ as the forecast for $y_{T+h}$.

2. Iterated autoregression (IAR). We estimate $y_t = \rho_0 + \Sigma_{j=1}^p \rho_j y_{t-j} + \varepsilon_t$ (we use $p = 4$). The $h$-period forecast is constructed by iterating the one-step forecast forward.

3. Direct forecast from autoregression (DAR). For each $h$, we estimate $y_{t+h} = \rho_0 + \Sigma_{j=1}^p \rho_j y_{t-j} + \varepsilon_t$ (we use $p = 4$). Each $h$-step forecast is a one-step forecast from the model for the appropriate $h$. The iterated AR forecast will asymptotically outperform the direct model if the AR(4) model is correctly specified, but the direct forecast may be more robust to misspecification, as discussed by Marcellino, Stock and Watson (2006).

4. Equal weighted averaging (EWA). We first estimate and forecast using $n$ simple models, each of the form $y_{t+h} = \rho_0 + \Sigma_{j=1}^p \rho_j y_{t-j} + \beta_i x_{it} + \varepsilon_{it}$ for $i = 1, \ldots n$ (we use $p = 4$). Letting $\hat{y}_{T+h}^i$ be the forecast of $y_{T+h}$ from the $i$th model, the EWA forecast is $n^{-1}\Sigma_{i=1}^n \hat{y}_{T+h}^i$. This method was first proposed by Bates and Granger (1969) and its empirical success is part of the folklore of forecasting. The method is considered by Stock and Watson (2003), among others.

5. Bayesian model averaging (BMA). In this method, described in more detail by Wright (2003), we take the $n$ models as in EWA and form a weighted average forecast motivated by the following Bayesian reasoning. Assign a prior in which each model is equally likely to be the true model. For the prior for the model parameters, we follow Fernandez, Ley

and Steel (2001). Write each model as $y_t = \lambda_i' w_{it} + \varepsilon_{it}$. Assume that $\varepsilon_{it} \sim N(0, \sigma^2)$ and assume that the prior for $\lambda_i$ conditional on $\sigma$ is $N(\bar{\lambda}, \phi(\sigma^2 \Sigma_{i=1}^n w_{it} w_{it}')^{-1})$ ; $\bar{\lambda}$ is described below. The (improper) marginal prior for $\sigma$ is proportional to $1/\sigma$. In this case, the posterior mean of $\lambda_i$ is $\tilde{\lambda}_i = \frac{\hat{\lambda}_i \phi}{1-\phi} + \frac{\bar{\lambda}\phi}{1-\phi}$ where $\hat{\lambda}_i$ denotes the OLS estimate of $\lambda_i$ while the probability that the $i^{th}$ model is the true model is

$$P(M_i) = (\tfrac{1}{1+\phi})^{k/2} [(Y - W_i\bar{\lambda})'(Y - W_i\bar{\lambda}) - \tfrac{\phi}{1+\phi}(Y - W_i\bar{\lambda})'W_i(W_i'W_i)^{-1}W_i'(Y - W_i\bar{\lambda})]^{-T/2}$$

where $Y = (y_1, y_2, ...y_T)'$ and $W_i = (w_{i1}, w_{i2}, ...w_{iT})'$. Accordingly, the proposed forecast is $\Sigma_{i=1}^n P(M_i)\tilde{\lambda}_i' w_{it}$. In our implementation, we will set $p = 4$, and let $\bar{\lambda} = (\bar{\rho}_0, \bar{\rho}_1, ...\bar{\rho}_p, \bar{\beta}_i)'$ where $\bar{\beta}_i = 0$ and $\bar{\rho}_0, \bar{\rho}_1, ...\bar{\rho}_p$ are obtained from fitting an autoregression to the pre-sample consisting of the data for 1947Q1 to 1960Q1 as recorded in the 1978Q4 vintage of the Federal Reserve Bank of Philadelphia real-time dataset.

The theoretical justification of this method relies on strictly exogenous regressors and iid errors—assumptions that are patently false in our application. Earlier work (Koop and Potter (2003) and Wright (2003)) shows that the method works well in cases like the one at hand, however, and we simply view BMA as a pragmatic shrinkage device.

6. Factor augmented autoregression (FAAR): For each $h$, we estimate $y_{t+h} = \rho_0 + \Sigma_{j=1}^p \rho_j y_{t-j} + \Sigma_{i=1}^m \gamma_i z_{it} + \varepsilon_t$ where $\{z_{it}\}_{i=1}^m$ are the first $m$ principal components of $\{x_{it}\}_{i=1}^n$. The predictors are first standardized to have mean zero and unit variance. We use $p = 4$, $m = 3$. The forecasts are then constructed as in the direct AR forecast.

7. Factor augmented vector autoregression (FAV): $\xi_t = \mu_0 + \Sigma_{j=1}^{\bar{p}} \mu_j \xi_{t-j} + \varepsilon_t$, where $\xi_t = (y_t, z_{1t}, z_{2t}, ...z_{mt})'$. We set $\bar{p} = 1$ and $m = 3$. The model can be estimated and iterated forward to provide a forecast of $y_{T+h}$. This method was proposed by Bernanke, Boivin and Eliasz (2005).

8. An integrated factor augmented VAR (IFAV). $\xi_t = \mu_0 + \Sigma_{j=1}^{\bar{p}} \mu_j \xi_{t-j} + \varepsilon_t$, where $\xi_t = (\Delta y_t, z_{1t}, z_{2t}, ...z_{mt})'$. We set $\bar{p} = 1$ and $m = 3$. The model can be estimated and

iterated forward to provide a forecast of $y_{T+h}$. This is the same as 4, except that a unit root is being imposed in the variable to be forecast.

9. A dynamic factor model (DF). This method, described in detail by Forni, Hallin, Lippi and Reichlin (2005), takes the vector of the stochastic standardized components of the data, $z_t$, (both the variable being forecast and the predictors), and assumes that is can be represented as $z_t = \chi_t + \zeta_t$ where $\chi_t = B(L)f_t$ is a vector of common components, $\zeta_t$ is a vector of idiosyncratic components, $f_t$ is a $q$x1 vector of dynamic factors that follow a stationary vector autoregression of order $s$, and the processes $\chi_t$ and $\zeta_t$ are mutually orthogonal at all leads and lags. The identifying assumption is that the eigenvalues of the spectral density matrix of $\zeta_t$ are uniformly bounded in the limit as the number of predictors gets large, whereas the eigenvalues of the spectral density matrix of $\chi_t$ are unbounded as the number of predictors gets large. The model can be written in the form $z_t = CF_t + \zeta_t$ where $F_t = (f_t', f_{t-1}'...f_{t-s}')'$ is a vector of $r = q(s+1)$ static factors. Given an estimator of the spectral density matrix of $z_t$, Forni, Hallin, Lippi and Reichlin show how to use dynamic principal components analysis to estimate the spectral density matrices of $\chi_t$ and $\zeta_t$ and hence to construct $h$-step ahead forecasts of $\chi_t$. Following Forni, Hallin, Lippi and Reichlin and also D'Agostino and Giannone (2006), we use this as the $h$-step ahead forecast for $z_t$, treating $\zeta_t$ as white noise for prediction purposes. The method differs from standard principal components analysis in that the dynamic restrictions on $F_t$ are imposed[6] and the different series are weighted by their respective signal-to-noise ratios in estimating the factors. We set $q = 3$, $r = 15$ and estimate the spectral density matrix of $z_t$ using a Bartlett window with truncation lag equal to the square root of the sample size.[7]

10. An unobserved component stochastic volatility model (UCSV). Stock and Watson (2007) find that an unobserved components model with stochastic volatility provides

---

[6]Concretely, the restriction is that the spectral density matrix of $F_t$ must have rank $q$.

[7]We are grateful to Mario Forni for providing us with the code for implementing this procedure.

good forecasts for inflation. The model is a univariate specification that $y_t = \tau_t + \eta_t^T$ and $\tau_t = \tau_{t-1} + \eta_t^P$ where $\eta_t^T$ is $iidN(0, \sigma_{T,t}^2)$, $\eta_t^P$ is $iidN(0, \sigma_{P,t}^2)$, $\log(\sigma_{T,t}^2) = \log(\sigma_{T,t-1}^2) + \psi_{1,t}$, $\log(\sigma_{P,t}^2) = \log(\sigma_{P,t-1}^2) + \psi_{2,t}$ and $(\psi_{1,t}, \psi_{2,t})'$ is $iidN(0, I_2)$. The interpretation of the model is that $y_t$ is the sum of an stochastic trend and noise, with both the volatility of the noise (temporary shocks) and the shocks to the stochastic trend (permanent shocks) being time-varying. The model implies that the series is I(1). Stock and Watson found that this model forecast inflation well because it recognized that the permanent component of inflation had high variance in the early 1980s, but became more stable subsequently. The model can be estimated by Markov Chain Monte Carlo. The forecast of $y_{T+h}$ from this model is simply the filtered estimate of $\tau_T$.

Finally, we note that throughout this paper we consider forecasting of one-quarter growth or inflation, $h$ quarters hence. Many authors instead consider the prediction of cumulative growth of inflation from quarter $t-1$ to quarter $t+h$, or four-quarter growth or inflation ending $h$ quarters hence. These are of course all different ways of expressing the same information. However, since one of our purposes in this paper is to assess the relative information content and Greenbook forecasts at different horizons, it seemed best to us to report results in terms of one-quarter growth at different horizons, as these other measures confound short- and longer-term predictive ability.

3.2 *Bootstrap p-values for non-nested comparisons*

One of our goals is to compare the forecast accuracy of the Greenbook with the time series methods. The appropriate construction of p-values for a comparison of root mean square prediction errors depends on whether the forecasts being compared are nested, or non-nested. We think it is appropriate to assume that the time series forecasts are not nested in the Greenbook. We accordingly approximate the sampling distribution of the RMSPEs using the moving-blocks bootstrap of Künsch (1989) and Liu and Singh (1992).[8]

---

[8] We use a block length of 10, corresponding to a span of a bit more than one year of data.

This is effectively using the moving-blocks bootstrap to simulate the distribution of the Diebold and Mariano (1995) statistic.[9]

We would also like to test the relative merits of the different atheoretical time series methods. But many pairs of forecasting models that we consider are nested. The test statistic of Diebold and Mariano (1995) has a nonstandard distribution under the null hypothesis of equal forecast accuracy if the models are nested, as the models are the same under the null (Clark and McCracken (2001)), and the bootstrap is not valid even to first order. Thus our bootstrap p-values do not allow us to make nested model comparisons.

## 4. Results

Our main metric for forecast accuracy is the root mean square prediction error (RMSPE) for the various forecasting methods (Table 3). We provide RMSPEs at each horizon 0 through 5—that is, the current quarter through 5 quarters hence. For each of the time series forecasting methods, we provide RMSPEs of forecasts that condition on data up to and including quarter $t - 1$, where quarter $t$ denotes the quarter in which the Greenbook was published. We also report the results of the time series forecasts that take later jumping-off points, i.e. that condition on data up to and including quarter $t$, $t + 1$, $t + 2$ or $t + 3$, using Greenbook forecasts for current and future quarters. Presuming that Greenbook is best at very short horizons, this method allows us to ask if there is some horizon at which one can usefully switch from Greenbook to one of the atheoretical methods.

When the RMSPE for a time series forecasting model is less than the corresponding value for Greenbook, the entry is in bold. Cases in which the RMSPE is significantly different from the corresponding value for Greenbook are marked at 1, 5 and 10 percent significance levels, using the two-tailed Diebold-Mariano test with bootstrap $p$-values as

---

[9]The asymptotic p-values for the Diebold-Mariano test using Newey-West standard errors with a lag length of 10 are not shown, but are similar to those implied by the bootstrap.

described above.

*4.1 Greenbook versus atheoretical models large and small*

The results for inflation and output growth are strikingly different. For inflation (either the deflator or CPI inflation), a quick summary goes like this. Greenbook dominates all the time series methods at all forecasting horizons and for almost all jumping-off points.[10] The point estimates for Greenbook are typically 10 to 40 percent smaller than those of the other models and we can reject the hypothesis of equal forecast accuracy of Greenbook and the other methods.

The main exception to this general result is that as we move the jumping-off point out in time, we cannot reject the hypothesis that Bayesian model averaging is as accurate as Greenbook. However, the point estimates for Greenbook remain smaller than those for BMA, except for the farthest out jumping-off points and horizons.

For output growth, Greenbook is a good deal better than the atheoretical methods at horizon zero. This is consistent with the view that the Fed usefully exploits a great deal more information about the current state of the economy than is used in the time series models. After horizon zero, however, the advantage largely evaporates. One can see this in two ways. First, if we keep the jumping-off point at $t-1$, the point estimates of Greenbook and the other models are comparable as we consider horizons beyond zero. Thus, the Greenbook advantage at time zero does not translate into forecasting gains at other horizons. Second, if we move the jumping-off point out even one quarter, the Greenbook advantage at all remaining horizons disappears, and is perhaps even reversed. That is, for several time series methods the point estimate of the RMSPE is smaller and we cannot reject the hypothesis that the method is at least as good as Greenbook.

As another metric for comparing the Greenbook and atheoretical forecasts, Table 4

---

[10]There appear to be quite strong seasonal patterns in many vintages of the GDP/GNP deflator inflation data, despite the fact that these data are seasonally adjusted. We experimented with adding deterministic seasonal dummies to each of the forecasting methods, but found that this nearly always increased mean-square prediction errors.

shows the proportion of forecasts for which each atheoretical model is more accurate *ex-post* than the Greenbook. Thus, numbers greater than 50 percent favor the atheoretical forecasts, while numbers smaller than 50 percent favor the Greenbook. For inflation, Greenbook does generally does better than the time series forecasts a good bit more than half the time. An exception is that BMA does better than Greenbook a bit more than half the time in forecasting CPI inflation at longer horizons. For growth, Greenbook does better than the time series methods well more than half the time at horizon zero, but at all other horizons Greenbook and time series models seem about equally likely to be more accurate.

These results suggest a significant difference in the forecastability of output and inflation. This result becomes even more distinct in the next section.

*4.2 Comparison of atheoretical methods*

We are also interested in evaluating the relative merits of large dataset methods versus univariate methods. For inflation (either deflator or CPI), among the univariate forecasts, the iterated autoregression and UCSV model generally seem to do best. However, Bayesian model averaging does a good bit better than any of the univariate inflation forecasts, and generally gives the smallest RMSPE among all the atheoretical inflation forecasts considered. The factor-augmented vector autoregression and dynamic factor forecasts are notably *less* accurate predictors of inflation than the iterated autoregression and UCSV forecasts. However, the factor augmented-vector autoregression that imposes a unit root in inflation (IFAV) does better than the univariate forecasts, though still not quite as well as BMA.

For growth, among the univariate forecasts, the iterated autoregression generally gives the smallest RMSPE, and the univariate forecasts that impose a unit root (random walk and UCSV) predict poorly. The forecasts based on equal-weighted averaging, Bayesian model averaging and the factor-augmented vector autoregression all typically

14

give smaller RMSPE than the iterated AR forecast, but the gains are modest. Even this advantage, however, largely disappears when we consider only the period since what has become known as the Great Moderation, as discussed below.

*4.3 Comments*

One striking observation from Table 3 is how small the Greenbook prediction errors for inflation are relative to any of the other forecasting procedures. This is consistent with Ang, Bekaert and Wei (2006) who find that private sector surveys outperform both univariate and multivariate time series forecasts of inflation. Ang, Bekaert and Wei make a conjecture similar to that of Sims (2002) regarding the Greenbook: the subjective methods may be able to aggregate very diverse information and to adapt rapidly to changes or special circumstances. These results sharpen those of Ang, Bekaert and Wei, however, in showing that the advantage of the subjective methods over atheoretical methods in forecasting inflation remains even when the atheoretical methods take advantage of the subjective assessment up through three quarters in the future.[11] Whether this advantage of Greenbook stems from access to a greater range of information or from a more sophisticated use of the available information, or both, remains an open question.

Our results show a clear pattern favoring the model averaging methods over the factor model approaches. Both model averaging approaches are consistently among the best in the results so far, with Bayesian model averaging perhaps having a slight edge over equal-weighted averaging. This result remains throughout the robustness checks performed below. Some factor model methods perform very badly in some cases and none of these methods consistently performs as well as the Bayesian model averaging approach. The fact that the factor model methods sometimes fare poorly was also noted by Bernanke and Boivin (2003). As they note, factor models fare better using the larger

---

[11]It would be natural to compare the Greenbook forecasts in this dataset with those from private sector forecasts. We leave this for future work, but note that this comparison is complicated by timing differences between the Greenbook and private sector forecasts.

dataset of Stock and Watson (2002). We take up this issue further below.

Notice further that while the Greenbook RMSPE profile increases with horizon for deflator inflation, the profile is essentially flat from horizon 1 onward for CPI growth and output growth. Thus, inflation measured by the GDP deflator becomes increasingly difficult to forecast as the horizon increases, and more data and subjective methods both seem to help. In contrast, GDP growth remains equally forecastable over horizons 1 through 5 and there are no clear gains from using more data or the subjective methods from Greenbook.

## 5. Additional topics

In this section, we take up several additional topics that shed light on the robustness of the results.

*5.1 Sub-samples and the Great Moderation*

The period from 1979–1983 was especially volatile in the U.S. economy, containing the sharpest and deepest recessions in the post-War era. The period since 1982 has seen the Great Moderation in which the economy seems much less volatile even than the period before 1979. Tulip (2005) has documented that the Greenbook forecast errors for output were largest early in the sample period.

To assess the importance of the early sample in our results, we construct results analogous to Table 3 for the period since 1984 (Table 5). As expected, the RMSPEs are substantially smaller than for the whole period in almost every case. Our qualitative conclusions, however, remain essentially the same. One conclusion is somewhat strengthened by excluding the turbulent period of the early 1980s: for forecasting growth, the iterated AR now performs even better compared to large dataset methods than before.

*5.2 Bias versus standard deviation*

In Table 6 we show the decomposition of the RMSPEs for the Greenbook forecasts and

statistical forecasts using data up to quarter $t-1$ into bias and standard deviation components. Results are shown for both the full sample and the period since 1984. For the inflation forecasts, all have some upward bias, at least at horizons beyond two quarters. The bias is worse for the stationary univariate and factor model time series forecasts than it is for the Greenbook forecasts, the random walk, IFAV and UCSV forecasts, and the predictions based on Bayesian model averaging. The standard deviations of the inflation forecasts can be thought of as bias-adjusted versions of the RMSPEs. As can be seen in Table 6, the standard deviations of the inflation forecasts are quite similar to the RMSPEs that we reported in Tables 3 and 5. The Greenbook inflation forecast errors have a smaller standard deviation than any of the statistical forecasts. Turning to the growth forecasts, the Greenbook forecasts have a modest downward bias, while all of the statistical models have an upward bias, except for the random walk and UCSV forecasts. The standard deviations of the growth forecasts are again very similar to the RMSPEs that we reported in Tables 3 and 5.

5.3 *Comparison with Atkeson and Ohanian*

In contrast to our results, Atkeson and Ohanian (2001) compared the RMSPE of Greenbook and random walk forecasts of inflation, and found that they were roughly equal, concluding that there was no incremental information in the Greenbook. The conclusion was based on 13 observations taken from the last Greenbook forecast in each year from 1983 to 1995 inclusive. They compared the Greenbook projection of GNP/GDP deflator inflation over the subsequent four quarters with a random walk forecast computed as the value of inflation over the previous four quarters. Thus, their measure of inflation and their version of the random walk forecast are different from ours.

We came very close to replicating their results by using their definitions and limiting our focus to the 13 observations they studied. We find that the RMSPE of the Greenbook forecast, relative to that of the random walk, was 0.96. Using their definitions but our

17

full sample, the ratio is more favorable to Greenbook at 0.73.

We examine the sampling properties of the estimate based on 13 observations using the same bootstrap procedure used above. Figure 1 shows the bootstrap approximation to the sampling distribution of the ratio of RMSPEs. It can clearly be seen that while this is centered around 1, in line with the Atkeson and Ohanian result, the ratio is quite imprecisely estimated. A 95 percent confidence interval for this ratio would span from 0.67 to 1.28, which includes the estimates for this ratio that we obtain with our much larger sample size. Thus, the good performance of the random walk 13 observations is entirely consistent with our finding that Greenbook dominates the random walk forecast.

5.4 *Real-time versus ex-post revised data*

All of the results presented so far use our real-time dataset. Most earlier work in the literature is based on a single recent vintage of data, and, hence, is not strictly comparable to our work. For comparison purposes, and to shed light on the importance of vintage issues, we repeat our exercise using a single vintage of data as observed in the Greenbook dated December 14, 2000 (the last vintage that is outside of the five-year window during which the data are confidential).

In this experiment, for each Greenbook, we took the Greenbook forecasts and compared these with the data as observed in the December 2000 vintage data. Then we constructed each of the statistical forecasts described above using observations for all the predictors from the December 2000 Greenbook at the quarterly frequency using a recursive out-of-sample scheme. For example, we used the December 2000 Greenbook vintage data from 1960Q2 through 1995Q3 to construct forecasts for each quarter from 1995Q4 through 1996Q4 (horizons 0 through 5). And we again compared these forecasts with the data as observed in the December 2000 Greenbook. Only forecasts for 2000Q3 and prior quarters are included, because data for subsequent quarters were not yet available in December 2000, and so would not be used by a researcher in a standard

forecast evaluation exercise.

The RMSPEs are shown in Table 7. The RMSPEs obtained using *ex-post* revised data are uniformly smaller than their counterparts using real-time data (Table 3). Also, the atheoretical forecasts perhaps fare a little better relative to Greenbook in the revised dataset than using real-time data. This is true for both the univariate forecasts and for the forecasts using a large number of predictors. The relative performance of the large-dataset methods and univariate forecasts is however comparable in the real-time and revised datasets.

Our dataset is large, but not as large as some datasets using many predictors, such as the datasets used in Stock and Watson (2002, 2005). Bernanke and Boivin (2003) found that factor models gave better forecasts when using the 215-variable database of Stock and Watson (2002) than when using their real-time database of 78 variables, or when using the *ex-post* revised observations on these 78 variables.

For comparison purposes, we re-did the analysis of the previous section but using the database of Stock and Watson (2005) which is, of course, revised data.[12] The RMSPEs are shown in Table 8. The FAAR, factor-augmented VAR and dynamic factor forecasts give considerably better forecasts of inflation when using this dataset than when using the final (December 2000) vintage of our Greenbook dataset. The same is true, to a much lesser extent, for the IFAV forecasts that impose a unit root. The results in Tables 7 and 8 are otherwise not very different, but, for inflation, factor forecasts seem to work best when using more highly disaggregated data than we have.

These results present the same puzzle as noted by Bernanke and Boivin (2003): the factor models perform much more poorly in forecasting inflation in fairly large real-time datasets than in the much the larger, fully-revised dataset of Stock and Watson. There are three natural hypotheses about this puzzle. First, it may be that the data revisions

---

[12] This dataset contains 135 variables and is larger than our database, but still smaller than the 215-variable database of Stock and Watson (2002) that does not however extend through 2000.

account for the difference. Second, it may be that the larger dataset simply has more relevant information than the smaller one. Third, it may be that while the forecast-relevant information may be similar in the large and small datasets, the factor models extract that information more effectively in the larger dataset.

With the available data, we cannot fully discriminate among these three hypotheses. The results of Table 7 and analogous results of Bernanke and Boivin suggest that the first hypothesis is not the main story since, with our datasets, the relative performance of the factor-based forecasts and the univariate time series forecasts is comparable in the real-time data (Table 3) and the revised data (Table 7).

Our results also suggest that the second hypothesis may not be a major factor. Despite the fact that the performance of the factor models varies considerably as we consider different datasets, the model averaging approaches give good performance and similar performance throughout. For example, the Table 7 and Table 8 results for the model averaging forecast are very similar. Thus, there is no strong support for the view that the larger dataset contains important information about output and inflation that is simply missing from the smaller dataset.

Turning to the third hypothesis, it may be that adding more variables—or adding the particular mix of variables that happen to be added by Stock and Watson (2002, 2005)—allows the factor methods more clearly to isolate the important information that the model averaging methods more robustly isolate across the various cases. Perhaps this should not be surprising. The factor model approaches merely pick out a few dominant factors from the large dataset, not focusing on the ability of those factors to forecast inflation or output directly. Loosely speaking, we never ask the full data the question we are ultimately interested in. One might suspect that such methods could be sensitive to which variables happen to be included.

In contrast, the model averaging methods directly exploit the information about each variable for the question of interest. But these methods do so in a very limited

way—we consider a large number of bivariate models. The the distinction between these methods is in part a distinction between bivariate and more richly multivariate methods.

Of course, the relative merits of these methods could be formalized and assessed as hypotheses about the underlying factor structure in macro data. We leave this analysis for future work.

### 5.6 *Combining atheoretical and Greenbook forecasts*

The atheoretical and Greenbook forecasts have comparable predictive performance for growth beyond the current quarter, though not for inflation. Meanwhile, since they are not perfectly correlated, this implies that a better forecast must be available by combining the Greenbook and atheoretical forecasts, at least for growth. Letting $\hat{y}_{T+h}^{GB}$ and $\hat{y}_{T+h}^{BMA}$ denote the Greenbook and BMA forecasts, respectively, we can define a combined forecast as

$$\hat{y}_{T+h}^{COMB} = \lambda \hat{y}_{T+h}^{GB} + (1 - \lambda)\hat{y}_{T+h}^{BMA}$$

For any fixed $\lambda$, we can then evaluate the real-time RMSPE of this forecast and plot it against $\lambda$. These are shown in Figures 2 and 3 for inflation and growth forecasts respectively. The jumping off point is $t - 1$, and results for three horizons are shown. For inflation, the optimal choice of $\lambda$ is close to 1, which amounts to putting all of the weight on the Greenbook forecast. For growth, however, the optimal choice of $\lambda$ is about 0.5 at horizons beyond the current quarter, meaning that substantial shrinkage of the Greenbook towards the BMA forecast (or vice-versa) would give better predictions than either forecast alone. The reductions in RMSPE due to this shrinkage are, however, very modest.

### 5.6 *The Greenbook as a conditional forecast*

The Greenbook projection is conditioned on a hypothetical path of policy that is counterfactual in the sense that it is not supposed to be a forecast of policy (see Faust and

21

Wright (2006)). Nearly all work on assessing the information content of central bank forecasts ignores their conditional nature and assesses the forecasts as though they were unconditional forecasts. We have done so here, thereby implicitly assessing their properties when viewed as though they were unconditional forecasts. However it should be noted that Greenbook could appear better—or worse—if we were to take proper account of its conditional nature. Faust and Wright propose a method for backing out an implied unconditional forecast based on comparing the hypothetical path of policy with the path implied by money market futures rates, under the assumption that the latter is the unconditional expectation of future interest rates. In future work, we intend to include this implied unconditional counterpart of the Greenbook forecast in the forecast evaluation exercise.

## 6. Conclusions

In this paper, we compare the forecast accuracy of three types of forecasting models: small time series models that are atheoretical from an economic perspective, large atheoretical time series models, and the Federal Reserve's Greenbook forecast. We focus on predicting inflation and output growth in a large real-time dataset.

The Greenbook forecast comparison is interesting for a few reasons. This forecast is generally thought to be at or near the frontier of best accuracy among real-world forecasts. Further, it is based on a much wider array of information than even our large time-series models, and it is subjective and heavily influenced by the staff's understanding of economics. Thus, to some extent, this comparison allows us to measure whether one version of a much richer forecasting framework helps much in forecasting.

Sims (2002) and others have conjectured that this richer framework of the Greenbook forecast may be of most value in evaluating the current state of the economy—one can take account of strikes, hurricanes, and the timing of holidays. Also some of the ability of the Greenbook to assess the current state of the economy surely owes to the

attention that Federal Reserve staff pay to incoming macroeconomic data releases. The case is particularly clear for CPI inflation as these data are released at the monthly frequency. Depending on the time within the quarter when the Greenbook forecast is made, one or even two months of CPI inflation for the current quarter will already have been released and will clearly be factored into the current-quarter CPI inflation forecast. Authors including Giannone, Reichlin and Small (2007) and Aruoba, Diebold and Scotti (2007) have recently proposed powerful methods for measuring the current state of the economy from the real-time monitoring of incoming data releases.

One naturally wonders whether the advantage of Greenbook relative to the atheoretical methods conisdered in this paper decays quickly as the forecasting horizon grows. We examine this question by checking whether one could improve forecast accuracy by using hybrid forecast that is the Greenbook forecast through some horizon such as 1 or 2 quarters out and then switching to a time series method.

For both output and inflation, we find that the Greenbook forecast is quite good compared with purely atheoretical methods. Beyond this, the results for output and inflation are strikingly different.

For output, once we give the time series models the Greenbook forecast for the current quarter at the time of the forecast, the atheoretical models perform as well or better than Greenbook. Indeed, after the current quarter nothing does much better than the AR(1) model. Given a rich evaluation of the current state of the economy, it is hard to beat nearly the simplest of time series models.

For inflation, Greenbook seems to have a considerable advantage over the atheoretical methods even at long horizons and even when the atheoretical methods jump off only after several quarters of Greenbook forecast. We find it surprising that Greenbook continues to dominate, even at long horizons and that the results for output and inflation are so different.

The results also shed some light on comparisons among the competing time series

methods. Two summary points are worth emphasizing. First, the results are generally less supportive of the large model methods than some earlier work. The large models offer no substantial gain for output growth prediction and only the best of the methods offer an advantage for inflation. Second, the large model forecasts that are built up using an average of bivariate models consistently outperform those that are more richly multivariate. That is, Bayesian and simple model averaging tend to dominate the models attempting to exploit some richer factor structure in the macro data. This result is consistent with earlier work using a single vintage and with earlier work on model averaging more generally. The results raise potentially important questions about the underlying factor structure of the macro data.

## References

Ang, A., G. Bekaert and M. Wei (2006): Do Macro Variables, Asset Markets or Surveys Forecast Inflation Better?, *Journal of Monetary Economics*, forthcoming.

Aruoba, S.B., F.X. Diebold and C. Scotti (2007): Real-Time Measurement of Business Conditions, working paper.

Atkeson, A. and L.E. Ohanian (2001): Are Phillips Curves Useful for Forecasting Inflation?, *Federal Reserve Bank of Minneapolis Quarterly Review*, 25, pp.2-11.

Bates, J.M. and C.W.J. Granger (1969): The Combination of Forecasts, *Operations Research Quarterly*, 20, pp.451-468.

Bernanke, B.S. and J. Boivin (2003): Monetary Policy in a Data-Rich Environment, *Journal of Monetary Economics*, 50, pp.525-546.

Bernanke, B.S., J. Boivin and P. Eliasz (2005): Measuring Monetary Policy: A Factor Augmented Vector Autoregressive (FAVAR) Approach, *Quarterly Journal of Economic*s, 120, pp.387-422.

Boivin, J. and S. Ng (2005): Monetary Policy in a Data-Rich Environment, *Journal of Monetary Economics*, 50, pp.525-546.

Clark, T.E. and M. McCracken (2001): Tests of Equal Forecast Accuracy and Encompassing for Nested Models, *Journal of Econometrics*, 105, pp.85-110.

D'Agostino, A. and D. Giannone (2006): Comparing Alternative Predictors Based on Large Panel Factor Models, European Central Bank Working Paper 680.

Diebold, F.X. and R.S. Mariano (1995): Comparing Predictive Accuracy, *Journal of Business and Economic Statistics,* 13, pp.253-263.

Edge, R.M., M. T. Kiley and J.P. Laforte (2006): A Comparison of Forecast Performance between Federal Reserve Staff Forecasts, Simple Reduced form Models and a DSGE Model, working paper.

Faust, Jon and Jonathan H. Wright (2006): Efficient Forecast Tests for Conditional Policy Forecasts, working paper.

Fernandez, C., E. Ley and M.F.J. Steel (2001): Model Uncertainty in Cross-Country Growth Regressions, *Journal of Applied Econometrics*, 16, pp.563-576.

Forni, M., M. Hallin, M. Lippi and L. Reichlin (2005): The Generalized Dynamic Factor Model: One-Sided Estimation and Forecasting, *Journal of the American Statistical Association*, 100, pp.830-840.

Giannone, D., L. Reichlin and L. Sala (2004): Monetary Policy in Real Time, in M. Gertler and K. Rogoff (eds.), NBER Macroeconomics Annual, 2004, MIT Press, Cambridge.

Giannone, D., L. Reichlin and D. Small (2007): Nowcasting GDP: The Real Time Informational Content of Macroeconomic Data Releases, working paper.

Koop, G. and S. Potter (2003): Forecasting in Large Macroeconomic Panels Using Bayesian Model Averaging, Federal Reserve Bank of New York Staff Report 163.

Künsch (1989): The Jackknife and the Bootstrap for General Stationary Observations, *Annals of Statistics*, 17, pp.1217-1241.

Liu, R. and K. Singh (1992): Moving Blocks Jackknife and Bootstrap Capture Weak Dependence, in *Exploring the Limits of the Bootstrap* (R. Lepage and L. Billard, eds.), Wiley, New York.

Marcellino, M., J.H. Stock and M.W. Watson (2006): A Comparison of Direct and Iterated Multistep AR methods for Forecasting Macroeconomic Time Series, *Journal of Econometrics*, 135, pp.499-526.

Reifschneider, D.L., D.J. Stockton and D.W. Wilcox (1997): Econometric Models and the Monetary Policy Process, *Carnegie Rochester Series on Public Policy*, 47, pp.1-37.

Romer, C.D. and D.H. Romer (2000): Federal Reserve Information and the Behavior of Interest Rates, *American Economic Review*, 90, pp.429-457.

Sims, C.A. (2002): The Role of Models and Probabilities in the Monetary Policy Process, *Brookings Papers on Economic Activity*, 2, pp.1-40.

Stock, J.H. and M.W. Watson (1999): Forecasting Inflation, *Journal of Monetary Economics*, 44, pp.293-335.

Stock, J.H. and M.W. Watson (2002): Forecasting Using Principal Components from a Large Number of Predictors, *Journal of the American Statistical Association*, 97, pp.1167-1179.

Stock, J.H. and M.W. Watson (2003): Forecasting Output and Inflation: The Role of Asset Prices, *Journal of Economic Literature*, 41, pp.788-829.

Stock, J.H. and M.W. Watson (2005): An Empirical Comparison of Methods for Forecasting Using Many Predictors, working paper.

Stock, J.H. and M.W. Watson (2007): Has Inflation Become Harder to Forecast?, *Journal of Money, Credit and Banking*, 39, pp.3-34.

Tetlow, R. and B. Ironside (2004): Real-Time Model Uncertainty in the United States: the Fed from 1996 to 2003, working paper.

Tulip, P. (2005b): Has Output Become More Predictable? Changes in Greenbook Forecast Accuracy, Finance and Economics Discussion Series 2005-31, Federal Reserve Board.

Wright, J.H. (2003): Forecasting U.S. Inflation by Bayesian Model Averaging, International Finance Discussion Paper 780, Federal Reserve Board.

Table 1: Dates of Greenbook Forecasts Employed

| Date | Date | Date | Date |
|------|------|------|------|
| 19800314 | 19850515 | 19901107 | 19961106 |
| 19800416 | 19850703 | 19901212 | 19961212 |
| 19800514 | 19850814 | 19910130 | 19970129 |
| 19800702 | 19850925 | 19910320 | 19970319 |
| 19800806 | 19851030 | 19910508 | 19970515 |
| 19800910 | 19851211 | 19910626 | 19970625 |
| 19801015 | 19860205 | 19910814 | 19970814 |
| 19801112 | 19860326 | 19910925 | 19970924 |
| 19801212 | 19860514 | 19911030 | 19971106 |
| 19810128 | 19860702 | 19920130 | 19971211 |
| 19810325 | 19860813 | 19920325 | 19980128 |
| 19810513 | 19860917 | 19920514 | 19980325 |
| 19810701 | 19861029 | 19920624 | 19980514 |
| 19810812 | 19861210 | 19920813 | 19980624 |
| 19810930 | 19870204 | 19920930 | 19980817 |
| 19811110 | 19870325 | 19921112 | 19980923 |
| 19811216 | 19870513 | 19921216 | 19981112 |
| 19820324 | 19870701 | 19930129 | 19981216 |
| 19820512 | 19870812 | 19930317 | 19990128 |
| 19820623 | 19870916 | 19930514 | 19990324 |
| 19820929 | 19871028 | 19930630 | 19990513 |
| 19821110 | 19871209 | 19930811 | 19990625 |
| 19821215 | 19880203 | 19930915 | 19991001 |
| 19830202 | 19880323 | 19931110 | 19991215 |
| 19830323 | 19880511 | 19931215 | 20000127 |
| 19830518 | 19880622 | 19940131 | 20000327 |
| 19830706 | 19881207 | 19940316 | 20000512 |
| 19830817 | 19890201 | 19940513 | 20000621 |
| 19830928 | 19890322 | 19940629 | 20000816 |
| 19831109 | 19890510 | 19940921 | 20000927 |
| 19831214 | 19890628 | 19941109 | 20001108 |
| 19840125 | 19890816 | 19941214 | 20001214 |
| 19840321 | 19890927 | 19950125 | |
| 19840516 | 19891108 | 19950322 | |
| 19840711 | 19891213 | 19950517 | |
| 19840815 | 19900131 | 19950628 | |
| 19840926 | 19900321 | 19950920 | |
| 19841031 | 19900509 | 19951108 | |
| 19841212 | 19900627 | 19951214 | |
| 19850206 | 19900815 | 19960626 | |
| 19850320 | 19900926 | 19960918 | |

This table lists the dates of publication of each Greenbook in our sample.

Table 2. Variables, transforms, and coverage

| Variable | Transform | First | Last | Number |
|---|---|---|---|---|
| Consumer spending on durable goods, real | DLN | 19800314 | 20001214 | 145 |
| Consumer spending on durable goods, nom | DLN | 19800314 | 19951214 | 130 |
| Exports, real | DLN | 19800314 | 20001214 | 145 |
| Exports, nom | DLN | 19800314 | 20001214 | 145 |
| Government spending, real | DLN | 19960626 | 20001214 | 34 |
| Government spending, nominal | DLN | 19800314 | 19951214 | 130 |
| Residential construction spending, real | DLN | 19800314 | 20001214 | 145 |
| Residential construction spending, nom | DLN | 19800314 | 20001214 | 140 |
| Inventory investment, nominal | FD | 19800314 | 20001214 | 145 |
| Inventory investment, real | FD | 19800314 | 20001214 | 145 |
| Inventory investment, manuf. and trade, real | FD | 19960626 | 19980624 | 16 |
| Imports, real | DLN | 19800314 | 20001214 | 145 |
| Imports, nom | DLN | 19800314 | 20001214 | 145 |
| Pesronal consumption spending, real | DLN | 19800314 | 20001214 | 145 |
| Pesronal consumption spending, nom | DLN | 19800314 | 20001214 | 145 |
| Spending on producers' durable equipment, real | DLN | 19800314 | 20001214 | 145 |
| Spending on producers' durable equipment, nom | DLN | 19800314 | 20001214 | 140 |
| Spending on producers' structures, real | DLN | 19800314 | 20001214 | 145 |
| Spending on producers' structures, nom | DLN | 19800314 | 20001214 | 140 |
| Merchandise exports | DLN | 19800314 | 19920514 | 103 |
| Foreign nom GNP index (bilateral weights) | DLN | 19821215 | 19911030 | 69 |
| Foreign short-term interest rate | Level | 19800314 | 19930129 | 109 |
| Stock of consumer durables, real | DLN | 19800314 | 20001214 | 145 |
| Stock of nonfarm inventories, real | DLN | 19800314 | 20001214 | 148 |
| Stock of nonfarm inventories, ex manuf. and trade | DLN | 19960626 | 19980624 | 16 |
| Stock of nonfarm nondurable inventories, real | DLN | 19800314 | 19880622 | 76 |
| Stock of nonfarm nonretail inventories, real | DLN | 19800314 | 19951214 | 130 |
| Stock of nonfarm retail durable inventories, real | DLN | 19800314 | 19951214 | 130 |
| Net stock of producers' durable equipment, real | DLN | 19800314 | 20001214 | 145 |
| Net stock of producers' structures, real | DLN | 19800314 | 20001214 | 145 |

Table 2. Variables, transforms, and coverage, cont.

| Variable | Transform | First | Last | Number |
|---|---|---|---|---|
| Nonfarm business employment | DLN | 19800314 | 20001214 | 145 |
| Employment of nonfarm proprietors | DLN | 19800314 | 19951214 | 130 |
| Hours of employees; nonfarm business sector | DLN | 19800314 | 20001214 | 145 |
| Unemployment | Level | 19800314 | 19951214 | 130 |
| M1 | DLN | 19800314 | 19951214 | 109 |
| M2 | DLN | 19800314 | 20001214 | 157 |
| Commercial and industrial loans at banks | DLN | 19800314 | 19921216 | 108 |
| Nonborrowed reserves | DLN | 19800314 | 19951214 | 130 |
| Currency plus travelers checks | DLN | 19800314 | 19951214 | 130 |
| Durable goods consumption deflator | DLN | 19800314 | 19951214 | 130 |
| Consumption deflator | DLN | 19800314 | 19951214 | 130 |
| GNP deflator | DLN | 19800314 | 19951214 | 130 |
| Average price per barrel of imported oil | DLN | 19800314 | 20001214 | 145 |
| Wholesale price index for fuels | DLN | 19800314 | 19951214 | 130 |
| Employee compensation per hour | DLN | 19800314 | 20001214 | 145 |
| Foreign exchange rate index, bilateral weights | DLN | 19800314 | 19911030 | 100 |
| Corporate bond rate; Moody's seasoned AAA | Level | 19800314 | 20001214 | 145 |
| Dividend-price ratio (Standard and Poors) | Log | 19800314 | 20001214 | 145 |
| Federal funds rate | Level | 19800314 | 20001214 | 145 |
| Ten-year Treasury CM yield | Level | 19800314 | 20001214 | 145 |
| 3 month Treasury bill rate | Level | 19800314 | 20001214 | 145 |
| Real GNP | DLN | 19800314 | 19951214 | 130 |
| Nominal GNP | DLN | 19800314 | 19951214 | 130 |
| Merchandise trade balance | FD | 19800314 | 19921112 | 107 |
| Net exports | FD | 19800314 | 19951214 | 130 |
| Compensation of employees | DLN | 19800314 | 19951214 | 130 |
| Corporate profits | DLN | 19800314 | 20001214 | 145 |
| Foreign output | DLN | 19790516 | 20001214 | 163 |
| Foreign consumer price index | DLN | 19790516 | 19930129 | 108 |
| Foreign exchange rate, multilateral weights | DLN | 19790516 | 19930129 | 108 |
| Hours; household and institutions sector | DLN | 19791114 | 19921112 | 98 |

Table 2. Variables, transforms, and coverage, cont.

| Variable | Transform | First | Last | Number |
|---|---|---|---|---|
| Total reserves | DLN | 19791114 | 19921216 | 102 |
| Civilian unemployment rate | Level | 19801112 | 19951214 | 92 |
| Stock of autos | DLN | 19820512 | 19890816 | 55 |
| Stock of consumer durables ex autos, real | DLN | 19820512 | 19951214 | 103 |
| Corporate bond rate | Level | 19820512 | 20001214 | 137 |
| Mortgage rate; effective annual yield | Level | 19820623 | 20001214 | 136 |
| Foreign real GNP index; bilateral weights | DLN | 19800314 | 19911030 | 100 |
| Output per hour | DLN | 19821215 | 19951214 | 99 |
| Final sales, real | DLN | 19840321 | 20001214 | 107 |
| Final sales, nominal | DLN | 19991001 | 20001214 | 10 |
| Spending on business fixed investment, real | DLN | 19841212 | 19951214 | 83 |
| Spending on business fixed investment, nom | DLN | 19841212 | 19951214 | 83 |
| Nominal exchange rate | DLN | 19850320 | 19930129 | 60 |
| Total cilivian employment | DLN | 19870204 | 19951214 | 66 |
| Stock of nonfarm nondurable inventories, real | DLN | 19881207 | 19951214 | 54 |
| Wholesale price of Petroleum products | DLN | 19890510 | 19951214 | 51 |
| Stock of motor vehicles and parts, real | DLN | 19890816 | 19951214 | 49 |
| Commodity price, industrial materials | DLN | 19900321 | 19901107 | 2 |
| Consumer Sentiment | Level | 19920130 | 19921112 | 7 |
| GDP Implicit Deflator | DLN | 19920130 | 20001214 | 64 |
| ReaL GDP | DLN | 19920130 | 20001214 | 64 |
| Nominal GDP | DLN | 19920130 | 20001214 | 64 |
| Current Account Balance | FD | 19920130 | 20001214 | 64 |
| Total Housing Stock | DLN | 19920514 | 20001214 | 62 |
| CPI | DLN | 19920514 | 19951214 | 28 |
| CPI ex food and energy | DLN | 19920514 | 19951214 | 28 |
| Producer Price Index | DLN | 19920624 | 19951214 | 27 |
| Real Disposable Income | DLN | 19920624 | 19951214 | 27 |
| Nominal Disposable Income | DLN | 19921216 | 19951214 | 23 |
| G10 real exchange rate | DLN | 19930915 | 20001214 | 51 |
| G18 real exchange rate | DLN | 19930317 | 19930811 | 4 |

Table 2. Variables, transforms, and coverage, cont.

| Variable | Transform | First | Last | Number |
|---|---|---|---|---|
| G18 nom exchange rate | DLN | 19930317 | 19930811 | 4 |
| Civilian employment | DLN | 19940316 | 19951214 | 13 |
| Civilian labor force | DLN | 19940316 | 20001214 | 47 |
| Price deflator for crude energy consumption | DLN | 19950517 | 20001214 | 39 |
| Compensation per hour | DLN | 19950517 | 19951214 | 5 |
| Consumption, energy sector | DLN | 19960626 | 20001214 | 34 |
| Import Price of Petroleum products | DLN | 19960626 | 20001214 | 34 |
| Effective federal funds rate | Level | 19960626 | 20001214 | 34 |
| Effective ten-year TreasuryCM yield | Level | 19960626 | 20001214 | 34 |
| Effective three-month Treasury bill yield | Level | 19960626 | 20001214 | 34 |
| Core PCE price index | DLN | 19970814 | 20001214 | 26 |
| Labor Force Participation Rate | Level | 19980817 | 20001214 | 18 |
| Moody's BAA corporate bond yield | Level | 19990513 | 20001214 | 12 |
| Final sales of GDP | DLN | 19991001 | 20001214 | 10 |
| Gross private domestic investment | DLN | 20000621 | 20001214 | 5 |
| Nonfarm business sector workweek | LN | 20000621 | 20001214 | 5 |
| Personal saving | DLN | 20000621 | 20001214 | 5 |

Notes: For transforms, DLN means first difference of the natural logarithm, FD means first difference, Log means natural logarithm, level means no transform. The columns labelled first and last give the date of the first and last vintage in which the variable is present, respectively. The column labelled num gives the number of vintages in which the variable is present.

Table 3a: RMSPE of deflator inflation forecasts, 1979-2000

| hor. | GB | RW | IAR | DAR | EWA | BMA | FAAR | FAV | IFAV | DF | UCSV |
|---|---|---|---|---|---|---|---|---|---|---|---|
| | | | | | jump off -1 | | | | | | |
| 0 | 0.81 | 1.31• | 1.09• | 1.09• | 1.07• | 1.06• | 1.18• | 1.28• | 1.24• | 1.30• | 1.12• |
| 1 | 0.84 | 1.50• | 1.26• | 1.26• | 1.22• | 1.17• | 1.23• | 1.58• | 1.42• | 1.63• | 1.24• |
| 2 | 0.91 | 1.27• | 1.16• | 1.17• | 1.13• | 1.11· | 1.22• | 1.65• | 1.17• | 1.61• | 1.18• |
| 3 | 1.03 | 1.29○ | 1.32• | 1.34• | 1.31• | 1.10 | 1.48• | 1.94• | 1.18 | 1.78• | 1.29○ |
| 4 | 1.06 | 1.40○ | 1.45• | 1.47• | 1.47• | 1.19· | 1.65• | 2.21• | 1.42○ | 1.90• | 1.39○ |
| 5 | 1.30 | 1.71○ | 1.72• | 1.78• | 1.75• | 1.39 | 1.84• | 2.43• | 1.65· | 2.15• | 1.66· |
| | | | | | jump off 0 | | | | | | |
| 1 | 0.84 | 1.16• | 1.16• | 1.16• | 1.15• | 1.13• | 1.26• | 1.40• | 1.22• | 1.39• | 1.14• |
| 2 | 0.91 | 1.26• | 1.14• | 1.13• | 1.11• | 1.08· | 1.21• | 1.64• | 1.17• | 1.45• | 1.18• |
| 3 | 1.03 | 1.27○ | 1.23• | 1.23• | 1.20• | 1.06 | 1.36• | 1.84• | 1.08 | 1.53• | 1.24○ |
| 4 | 1.06 | 1.41· | 1.41• | 1.43• | 1.43• | 1.15 | 1.60• | 2.09• | 1.27 | 1.78• | 1.39· |
| 5 | 1.30 | 1.65 | 1.69• | 1.72• | 1.71• | 1.37 | 1.72○ | 2.43• | 1.54 | 2.13• | 1.64 |
| | | | | | jump off 1 | | | | | | |
| 2 | 0.91 | 1.19○ | 1.08• | 1.08• | 1.08• | 1.08• | 1.18• | 1.40• | 1.13• | 1.32• | 1.15• |
| 3 | 1.03 | 1.43○ | 1.26• | 1.26• | 1.24• | 1.05 | 1.30• | 1.68• | 1.28○ | 1.45• | 1.31○ |
| 4 | 1.06 | 1.41○ | 1.34• | 1.35• | 1.34• | 1.11 | 1.44• | 1.87• | 1.26○ | 1.58• | 1.36○ |
| 5 | 1.30 | 1.58 | 1.58○ | 1.62○ | 1.61○ | **1.28** | 1.71○ | 2.22• | 1.43 | 1.90• | 1.57 |
| | | | | | jump off 2 | | | | | | |
| 3 | 1.03 | 1.16· | 1.21• | 1.21• | 1.21• | 1.11 | 1.34• | 1.49• | 1.20○ | 1.36• | 1.21○ |
| 4 | 1.06 | 1.27• | 1.31• | 1.33• | 1.31• | 1.11 | 1.40• | 1.70• | 1.19○ | 1.40• | 1.29○ |
| 5 | 1.30 | 1.44 | 1.53• | 1.54• | 1.53• | **1.26** | 1.63• | 2.08• | 1.31 | 1.68• | 1.48 |
| | | | | | jump off 3 | | | | | | |
| 4 | 1.06 | 1.18· | 1.23• | 1.23• | 1.23• | 1.08 | 1.34• | 1.47• | 1.14○ | 1.32• | 1.22○ |
| 5 | 1.30 | 1.44 | 1.48• | 1.47• | 1.46○ | **1.23** | 1.57• | 1.90• | 1.32 | 1.56• | 1.45 |

Notes: Each column, 2–11, reports the root mean square prediction error (RMSPE) for the forecast labelled at top for the current quarter (hor= 0) through 5 quarters into the future. For forecasts other than Greenbook, the data on which the forecast is based is the real-time data augmented to bring the data up to the jumping off point using the Greenbook forecast. Bold indicates that the alternative forecast has smaller RMSPE than Greenbook; •,○,· indicate that the difference between the RMSPE for the model and Greenbook is statistically significant at the 1, 5, or 10 percent level, respectively, based on the bootstrap $p$-values for the DM test as described in the text.

Table 3b: RMSPE of CPI inflation forecasts, 1979-2000

| hor. | GB | RW | IAR | DAR | EWA | BMA | FAAR | FAV | IFAV | DF | UCSV |
|---|---|---|---|---|---|---|---|---|---|---|---|
| | | | | | | jump off -1 | | | | | |
| 0 | 0.82 | 2.05• | 1.75• | 1.75• | 1.72• | 1.62• | 1.66• | 1.85• | 1.93• | 1.76• | 1.92• |
| 1 | 1.66 | 2.20• | 1.80 | 1.82 | 1.78 | **1.56** | 1.89 | 2.15• | 2.06• | 1.94· | 1.92○ |
| 2 | 1.78 | 1.95 | 1.85 | 1.95 | 1.91 | **1.58** | 1.95 | 2.11○ | **1.73** | 1.94 | 1.93 |
| 3 | 1.83 | 2.51○ | 2.35• | 2.37• | 2.31• | 1.86 | 2.28○ | 2.61• | 2.26○ | 2.12 | 2.42○ |
| 4 | 1.60 | 2.57• | 2.23• | 2.22• | 2.19• | 1.85 | 2.50• | 2.77• | 2.51• | 2.29• | 2.36○ |
| 5 | 1.63 | 2.44○ | 2.25• | 2.20• | 2.12• | 2.21 | 2.43• | 2.70• | 2.35• | 2.19○ | 2.38○ |
| | | | | | | jump off 0 | | | | | |
| 1 | 1.66 | 2.19• | 1.80· | 1.80· | 1.78 | **1.65** | 1.75 | 2.04• | 2.11• | 1.73 | 1.93• |
| 2 | 1.78 | 2.23○ | 1.82 | 1.82 | 1.79 | **1.55** | 1.87 | 2.24• | 2.02· | **1.71** | 1.96· |
| 3 | 1.83 | 2.13 | 2.01 | 2.04 | 2.00 | **1.63** | 2.01 | 2.38· | 1.95 | 1.85 | 2.06 |
| 4 | 1.60 | 2.52○ | 2.32• | 2.41• | 2.37• | 1.90 | 2.41• | 2.65• | 2.45○ | 2.17• | 2.33○ |
| 5 | 1.63 | 2.61○ | 2.23• | 2.26• | 2.19• | 1.71 | 2.33• | 2.86• | 2.47• | 2.17• | 2.40○ |
| | | | | | | jump off 1 | | | | | |
| 2 | 1.78 | 1.98○ | 1.79 | 1.79 | 1.78 | **1.60** | **1.77** | 2.01• | 1.92 | 1.79 | 1.89 |
| 3 | 1.83 | 2.11 | 2.05 | 2.08 | 2.04 | **1.68** | 1.99 | 2.27· | 1.90 | 1.86 | 2.04 |
| 4 | 1.60 | 2.21 | 2.31• | 2.46• | 2.43• | 1.97 | 2.38○ | 2.34• | 2.10 | 1.90○ | 2.18 |
| 5 | 1.63 | 2.40○ | 2.27• | 2.30• | 2.27• | 1.74 | 2.22• | 2.64• | 2.26○ | 2.06• | 2.35○ |
| | | | | | | jump off 2 | | | | | |
| 3 | 1.83 | 1.94 | 1.92 | 1.92 | 1.90 | **1.70** | 1.86 | 2.10 | 1.90 | 1.93 | 1.93 |
| 4 | 1.60 | 1.86 | 2.07○ | 2.13• | 2.10○ | 1.76 | 2.08· | 2.14• | 1.81 | 1.81• | 1.92 |
| 5 | 1.63 | 2.16 | 2.24• | 2.33• | 2.30• | 1.85 | 2.18· | 2.48• | 2.12 | 1.84○ | 2.18· |
| | | | | | | jump off 3 | | | | | |
| 4 | 1.60 | 1.70· | 1.77• | 1.77• | 1.75○ | 1.62 | 1.76 | 1.86• | 1.66 | 1.77○ | 1.74 |
| 5 | 1.63 | 1.88· | 2.02• | 2.06• | 2.02• | 1.70 | 2.09○ | 2.18• | 1.86 | 1.80· | 1.95· |

Notes: Each column, 2–11, reports the root mean square prediction error (RMSPE) for the forecast labelled at top for the current quarter (hor= 0) through 5 quarters into the future. For forecasts other than Greenbook, the data on which the forecast is based is the real-time data augmented to bring the data up to the jumping off point using the Greenbook forecast. Bold indicates that the alternative forecast has smaller RMSPE than Greenbook; •,○,· indicate that the difference between the RMSPE for the model and Greenbook is statistically significant at the 1, 5, or 10 percent level, respectively, based on the bootstrap $p$-values for the DM test as described in the text.

Table 3c: RMSPE of output growth forecasts, 1979-2000

| hor. | GB | RW | IAR | DAR | EWA | BMA | FAAR | FAV | IFAV | DF | UCSV |
|---|---|---|---|---|---|---|---|---|---|---|---|
| | | | | | jump off -1 | | | | | | |
| 0 | 2.17 | 3.31∘ | 2.77 | 2.77 | 2.71 | 2.73 | 2.75∘ | 2.73∘ | 2.73· | 2.75∘ | 2.74 |
| 1 | 2.75 | 3.56 | 2.75 | 2.76 | **2.64** | **2.49** | 2.80 | **2.63** | 2.81 | **2.67** | **2.72** |
| 2 | 2.72 | 3.82∘ | 2.83 | 2.90 | 2.80 | 2.86 | 2.88 | **2.69** | 2.97 | 2.92 | 2.96 |
| 3 | 2.76 | 3.40∘ | **2.61** | **2.58** | **2.54** | **2.66** | 2.83 | **2.74** | 3.09· | **2.63** | 2.84 |
| 4 | 2.57 | 4.08● | 2.69 | 2.71 | 2.62 | 2.59 | **2.56** | 2.60 | 3.11 | 2.67 | 2.97· |
| 5 | 2.41 | 3.62● | 2.66 | 2.74 | 2.71 | 2.86 | 2.58 | 2.60 | 3.62∘ | 2.76 | 2.92∘ |
| | | | | | jump off 0 | | | | | | |
| 1 | 2.75 | 2.90 | **2.66** | **2.66** | **2.57** | **2.48** | **2.59** | **2.51** | 2.76 | **2.50** | **2.61** |
| 2 | 2.72 | 3.20 | 2.76 | 2.82 | **2.71** | **2.53** | 2.74 | **2.58** | 2.79 | 2.86· | 2.86 |
| 3 | 2.76 | 3.08· | **2.61** | **2.63** | **2.58** | **2.69** | 2.76 | **2.69** | **2.73** | **2.66** | **2.72** |
| 4 | 2.57 | 3.37● | 2.66 | 2.66 | 2.58 | 2.58 | **2.52** | 2.57 | 2.94 | 2.65 | 2.89· |
| 5 | 2.41 | 3.19● | 2.64 | 2.74 | 2.65 | 2.61 | 2.54 | 2.53 | 2.91· | 2.83· | 2.79∘ |
| | | | | | jump off 1 | | | | | | |
| 2 | 2.72 | 3.07 | 2.74 | 2.74 | **2.68** | **2.68** | **2.66** | **2.61** | **2.63** | 2.72 | 2.79 |
| 3 | 2.76 | 3.04 | **2.62** | **2.61** | **2.58** | **2.64** | **2.75** | **2.64** | 2.82 | **2.69** | **2.75** |
| 4 | 2.57 | 2.83· | 2.57 | 2.57 | **2.51** | **2.50** | 2.58 | 2.60 | 2.68 | 2.63 | 2.71 |
| 5 | 2.41 | 2.74 | 2.62 | 2.62 | 2.56 | 2.56 | 2.57 | 2.53 | 2.79 | 2.71· | 2.62 |
| | | | | | jump off 2 | | | | | | |
| 3 | 2.76 | 2.91 | **2.71** | **2.71** | **2.66** | **2.58** | **2.62** | **2.56** | 2.80 | **2.60** | 2.77 |
| 4 | 2.57 | 2.62 | **2.54** | **2.51** | **2.46** | **2.44** | **2.53** | **2.55** | 2.64 | 2.57 | 2.62 |
| 5 | 2.41 | 2.48 | 2.59 | 2.66 | 2.59 | 2.50 | 2.61 | 2.52 | 2.73 | 2.65∘ | 2.53 |
| | | | | | jump off 3 | | | | | | |
| 4 | 2.57 | 2.68 | **2.55** | **2.55** | **2.49** | **2.43** | **2.44** | **2.44** | 2.59 | **2.52** | 2.61 |
| 5 | 2.41 | 2.54 | 2.54 | 2.59 | 2.52 | **2.33** | 2.52 | 2.44 | 2.66 | 2.52 | 2.47 |

Notes: Each column, 2–11, reports the root mean square prediction error (RMSPE) for the forecast labelled at top for the current quarter (hor= 0) through 5 quarters into the future. For forecasts other than Greenbook, the data on which the forecast is based is the real-time data augmented to bring the data up to the jumping off point using the Greenbook forecast. Bold indicates that the alternative forecast has smaller RMSPE than Greenbook; ●,∘,· indicate that the difference between the RMSPE for the model and Greenbook is statistically significant at the 1, 5, or 10 percent level, respectively, based on the bootstrap p-values for the DM test as described in the text.

Table 4a: Percentage of time series deflator inflation forecasts that are more accurate than GB, 1979-2000

| hor. | RW | IAR | DAR | EWA | BMA | FAAR | FAV | IFAV | DF | UCSV |
|------|----|-----|-----|-----|-----|------|-----|------|----|------|
| jump off -1 | | | | | | | | | | |
| 0 | 39 | 34 | 34 | 30 | 32 | 30 | 27 | 37 | 33 | 35 |
| 1 | 27 | 25 | 29 | 32 | 32 | 33 | 27 | 30 | 21 | 34 |
| 2 | 41 | 36 | 34 | 38 | 37 | 41 | 17 | 39 | 26 | 37 |
| 3 | 42 | 37 | 33 | 35 | 39 | 37 | 22 | 47 | 22 | 34 |
| 4 | 35 | 27 | 26 | 26 | 37 | 28 | 14 | 40 | 17 | 41 |
| 5 | 38 | 23 | 19 | 21 | 47 | 30 | 16 | 37 | 19 | 41 |
| jump off 0 | | | | | | | | | | |
| 1 | 32 | 31 | 31 | 31 | 32 | 30 | 33 | 31 | 28 | 32 |
| 2 | 34 | 34 | 34 | 33 | 34 | 38 | 23 | 35 | 22 | 37 |
| 3 | 42 | 38 | 34 | 36 | 39 | 32 | 26 | 44 | 24 | 39 |
| 4 | 34 | 29 | 30 | 27 | 39 | 34 | 18 | 42 | 23 | 40 |
| 5 | 35 | 28 | 26 | 23 | 43 | 32 | 16 | 41 | 18 | 45 |
| jump off 1 | | | | | | | | | | |
| 2 | 38 | 36 | 36 | 37 | 34 | 40 | 32 | 41 | 22 | 39 |
| 3 | 35 | 33 | 32 | 34 | 39 | 33 | 24 | 37 | 25 | 40 |
| 4 | 36 | 31 | 31 | 30 | 37 | 31 | 21 | 37 | 22 | 39 |
| 5 | 36 | 36 | 32 | 32 | 44 | 31 | 19 | 39 | 23 | 36 |
| jump off 2 | | | | | | | | | | |
| 3 | 40 | 36 | 36 | 38 | 38 | 39 | 26 | 39 | 26 | 39 |
| 4 | 32 | 32 | 32 | 32 | 36 | 32 | 28 | 34 | 27 | 30 |
| 5 | 37 | 31 | 31 | 29 | 37 | 30 | 23 | 39 | 24 | 41 |
| jump off 3 | | | | | | | | | | |
| 4 | 39 | 34 | 34 | 35 | 39 | 33 | 28 | 39 | 30 | 37 |
| 5 | 41 | 33 | 32 | 31 | 40 | 30 | 24 | 42 | 28 | 37 |

Notes: Each column, 2–10, reports the share of forecast periods in which the GB forecast error is larger, *ex post*, in absolute value than the forecast error from the model labelled at the top. Entries greater than 50 percent indicate that the model forecast does better than GB in more than half of forecast periods and are in bold. See also notes to Table 3.

Table 4b: Percentage of time series CPI inflation forecasts that are more accurate than GB, 1979-2000

| hor. | RW | IAR | DAR | EWA | BMA | FAAR | FAV | IFAV | DF | UCSV |
|---|---|---|---|---|---|---|---|---|---|---|
| | | | | | jump off -1 | | | | | |
| 0 | 15 | 14 | 14 | 16 | 22 | 17 | 16 | 13 | 17 | 17 |
| 1 | 42 | 45 | 45 | 45 | 49 | 40 | 35 | 43 | 41 | 43 |
| 2 | 44 | 43 | 42 | 43 | 46 | 43 | 39 | 50 | 36 | 45 |
| 3 | 41 | 37 | 33 | 34 | **52** | 33 | 33 | 31 | 34 | 37 |
| 4 | 42 | 36 | 37 | 39 | **54** | 35 | 28 | 39 | 28 | 43 |
| 5 | 44 | 30 | 33 | 33 | **55** | 38 | 28 | 36 | 23 | 42 |
| | | | | | jump off 0 | | | | | |
| 1 | 37 | 39 | 39 | 39 | 44 | 44 | 39 | 38 | 40 | 48 |
| 2 | 45 | **51** | 50 | **52** | **56** | 46 | 37 | 46 | 43 | **52** |
| 3 | 48 | 47 | 45 | 46 | **56** | 43 | 43 | 48 | 46 | 48 |
| 4 | 37 | 37 | 34 | 34 | **54** | 39 | 25 | 36 | 30 | 33 |
| 5 | 44 | 34 | 32 | 35 | **61** | 37 | 31 | 43 | 26 | 37 |
| | | | | | jump off 1 | | | | | |
| 2 | 42 | 48 | 48 | 48 | **61** | **52** | 40 | **50** | 45 | 49 |
| 3 | 46 | 47 | 45 | 48 | **58** | 45 | 39 | 48 | 44 | **50** |
| 4 | 49 | 39 | 38 | 41 | **53** | 46 | 34 | 48 | 38 | 45 |
| 5 | 39 | 32 | 30 | 29 | **54** | 39 | 23 | 42 | 28 | 33 |
| | | | | | jump off 2 | | | | | |
| 3 | 46 | **51** | **51** | **53** | **63** | **54** | 42 | **55** | 42 | **52** |
| 4 | **52** | 41 | 41 | 41 | **59** | 46 | 40 | **51** | 43 | 49 |
| 5 | 47 | 32 | 34 | 38 | **59** | 44 | 31 | 46 | 41 | 39 |
| | | | | | jump off 3 | | | | | |
| 4 | 50 | 48 | 48 | **51** | **53** | 50 | 43 | **55** | 48 | 46 |
| 5 | 43 | 38 | 36 | 41 | **59** | 46 | 37 | 50 | 47 | 38 |

Notes: Each column, 2–10, reports the share of forecast periods in which the GB forecast error is larger, *ex post*, in absolute value than the forecast error from the model labelled at the top. Entries greater than 50 percent indicate that the model forecast does better than GB in more than half of forecast periods and are in bold. See also notes to Table 3.

Table 4c: Percentage of time series output growth forecasts that are more accurate than GB, 1979-2000

| hor. | RW | IAR | DAR | EWA | BMA | FAAR | FAV | IFAV | DF | UCSV |
|------|-----|-----|-----|-----|-----|------|-----|------|-----|------|
| jump off -1 | | | | | | | | | | |
| 0 | 25 | 37 | 37 | 39 | 41 | 38 | 34 | 36 | 37 | 31 |
| 1 | 39 | **54** | **57** | **51** | 48 | 41 | 43 | 43 | 46 | 46 |
| 2 | 39 | **51** | **52** | **54** | **52** | 45 | 47 | 45 | **50** | 39 |
| 3 | 45 | **54** | **55** | **53** | **50** | 47 | 43 | 41 | **52** | **50** |
| 4 | 36 | **50** | 44 | 45 | 45 | 40 | 43 | 41 | 44 | 40 |
| 5 | 36 | 41 | 43 | 46 | 40 | 43 | 43 | 38 | 39 | 37 |
| jump off 0 | | | | | | | | | | |
| 1 | **52** | **56** | **56** | **55** | **54** | 46 | 48 | 46 | 43 | 46 |
| 2 | **51** | **52** | **52** | **53** | **54** | 46 | 48 | **52** | 43 | 44 |
| 3 | 46 | **54** | **56** | **54** | **52** | 41 | 43 | 46 | 50 | **50** |
| 4 | 39 | **50** | **52** | **52** | 47 | 44 | 42 | 42 | 43 | 39 |
| 5 | 33 | 40 | 40 | 45 | 41 | 42 | 43 | 41 | 45 | 37 |
| jump off 1 | | | | | | | | | | |
| 2 | 49 | **52** | **52** | **54** | **57** | 49 | 48 | **52** | **52** | 45 |
| 3 | 48 | **54** | **56** | **53** | 49 | 43 | 44 | 47 | 42 | 48 |
| 4 | 45 | **50** | 49 | 49 | 48 | 48 | 45 | 42 | 43 | 43 |
| 5 | 40 | 42 | 43 | 44 | 41 | 44 | 43 | 39 | 43 | 41 |
| jump off 2 | | | | | | | | | | |
| 3 | 47 | **51** | **51** | **50** | **58** | 45 | 47 | 43 | **50** | **55** |
| 4 | **51** | 50 | **51** | **50** | **54** | 46 | 48 | 46 | 47 | 50 |
| 5 | 45 | 43 | 46 | 43 | 41 | 39 | 43 | 41 | 36 | 46 |
| jump off 3 | | | | | | | | | | |
| 4 | **51** | **53** | **53** | **56** | **59** | 50 | 47 | 43 | 47 | **55** |
| 5 | 48 | 46 | 47 | 44 | 45 | 39 | 44 | 41 | 45 | 48 |

Notes: Each column, 2–10, reports the share of forecast periods in which the GB forecast error is larger, *ex post*, in absolute value than the forecast error from the model labelled at the top. Entries greater than 50 percent indicate that the model forecast does better than GB in more than half of forecast periods and are in bold. See also notes to Table 3.

Table 5a: RMSPE of deflator inflation forecasts, 1984-2000

| hor. | GB | RW | IAR | DAR | EWA | BMA | FAAR | FAV | IFAV | DF | UCSV |
|------|------|------|------|------|------|------|------|------|------|------|------|
| | | | | | jump off -1 | | | | | | |
| 0 | 0.69 | 1.13• | 0.96• | 0.96• | 0.95• | 0.93• | 1.00• | 1.11• | 1.80• | 1.12• | 0.98• |
| 1 | 0.79 | 1.31• | 1.12• | 1.12• | 1.08• | 0.96• | 1.04• | 1.35• | 1.19• | 1.46• | 1.09• |
| 2 | 0.81 | 1.23• | 1.06• | 1.05• | 1.01• | 0.93• | 0.96∘ | 1.36• | 1.50• | 1.55• | 1.01• |
| 3 | 0.93 | 1.08∘ | 1.15• | 1.16• | 1.11• | 0.96∘ | 1.15• | 1.59• | 0.97 | 1.71• | 1.04∘ |
| 4 | 0.89 | 1.16∘ | 1.26• | 1.27• | 1.21• | 0.96· | 1.14• | 1.66• | 1.70∘ | 1.88• | 1.07∘ |
| 5 | 1.14 | 1.46∘ | 1.50• | 1.54• | 1.46• | **1.13** | 1.38• | 1.93• | 1.31 | 2.80• | 1.28 |
| | | | | | jump off 0 | | | | | | |
| 1 | 0.79 | 0.94∘ | 1.00• | 1.00• | 0.99• | 0.96• | 1.04• | 1.18• | 0.99• | 1.24• | 0.96∘ |
| 2 | 0.81 | 1.06• | 1.03• | 1.03• | 1.00• | 0.95• | 0.96∘ | 1.30• | 1.01• | 1.37• | 0.97∘ |
| 3 | 0.93 | 1.13• | 1.14• | 1.13• | 1.09• | 1.00· | 1.12• | 1.50• | 0.99 | 1.55• | 1.04∘ |
| 4 | 0.89 | 0.96 | 1.16• | 1.18• | 1.12• | 0.94 | 1.12• | 1.56• | **0.88** | 1.70• | 0.97 |
| 5 | 1.14 | 1.21 | 1.43• | 1.45• | 1.39• | **1.10** | 1.27 | 1.83• | 1.17 | 2.10• | 1.19 |
| | | | | | jump off 1 | | | | | | |
| 2 | 0.81 | 0.93∘ | 0.95• | 0.95• | 0.94• | 0.94• | 0.94∘ | 1.08• | 0.93∘ | 1.16• | 0.92∘ |
| 3 | 0.93 | 1.14• | 1.11• | 1.11• | 1.09• | 1.00 | 1.07• | 1.37• | 1.12• | 1.39• | 1.06∘ |
| 4 | 0.89 | 1.13• | 1.15• | 1.16• | 1.12• | 1.01· | 1.09• | 1.47• | 1.04∘ | 1.53• | 1.03∘ |
| 5 | 1.14 | 1.16 | 1.34∘ | 1.35∘ | 1.31· | **1.10** | 1.28· | 1.72• | **1.12** | 1.82• | 1.16 |
| | | | | | jump off 2 | | | | | | |
| 3 | 0.93 | 0.98 | 1.06• | 1.06• | 1.04• | 1.02∘ | 1.06• | 1.19• | 1.01· | 1.19• | 1.00· |
| 4 | 0.89 | 1.06∘ | 1.11• | 1.13• | 1.10• | 1.03∘ | 1.08• | 1.33• | 1.03∘ | 1.32• | 1.20∘ |
| 5 | 1.14 | 1.21 | 1.34∘ | 1.35∘ | 1.31∘ | 1.15 | 1.29∘ | 1.60• | **1.13** | 1.58• | 1.17 |
| | | | | | jump off 3 | | | | | | |
| 4 | 0.89 | 0.96 | 1.05• | 1.05• | 1.03• | 0.99∘ | 1.03• | 1.15• | 0.97· | 1.15• | 0.98· |
| 5 | 1.14 | 1.22 | 1.30∘ | 1.30∘ | 1.28∘ | 1.15 | 1.27∘ | 1.48• | 1.18 | 1.41• | 1.19 |

Notes: Each column, 2–11, reports the root mean square prediction error (RMSPE) for forecast labelled at top for the current quarter (hor= 0) through 5 quarters into the future. For forecasts other than Greenbook, the data on which the forecast is based is the real-time data augmented to bring the data up to the jumping off point using the Greenbook forecast. Bold indicates that the alternative forecast has smaller RMSPE than Greenbook; •,∘,· indicate that the difference between the RMSPE for the model and Greenbook is statistically significant at the 1, 5, or 10 percent level, respectively, based on the bootstrap $p$-values for the DM test as described in the text.

Table 5b: RMSPE of CPI inflation forecasts, 1984-2000

| hor. | GB | RW | IAR | DAR | EWA | BMA | FAAR | FAV | IFAV | DF | UCSV |
|------|------|------|------|------|------|------|------|------|------|------|------|
| | | | | | jump off -1 | | | | | | |
| 0 | 0.49 | 1.61• | 1.47• | 1.47• | 1.45• | 1.35• | 1.46• | 1.52• | 1.47• | 1.45• | 1.43• |
| 1 | 1.15 | 1.82• | 1.48· | 1.51∘ | 1.46· | 1.35 | 1.43 | 1.72• | 1.57• | 1.80• | 1.57• |
| 2 | 1.27 | 1.70 | 1.61 | 1.69 | 1.65 | 1.41 | 1.59 | 1.65 | 1.56 | 1.68∘ | 1.57 |
| 3 | 1.56 | 2.07· | 1.97∘ | 2.08∘ | 2.01∘ | 1.68 | 1.92· | 1.95· | 1.93∘ | 1.91∘ | 1.86 |
| 4 | 1.42 | 2.06• | 1.85• | 1.99• | 1.90• | 1.48 | 1.87• | 2.02• | 1.86• | 2.03• | 1.80∘ |
| 5 | 1.41 | 1.93∘ | 1.87• | 1.98• | 1.91• | 1.44 | 1.90• | 1.98• | 1.92• | 2.09• | 1.76· |
| | | | | | jump off 0 | | | | | | |
| 1 | 1.15 | 1.55• | 1.32∘ | 1.32∘ | 1.29∘ | 1.25· | 1.26· | 1.48• | 1.48• | 1.47∘ | 1.38• |
| 2 | 1.27 | 1.58∘ | 1.49 | 1.51 | 1.47 | 1.31 | 1.54 | 1.56∘ | 1.44· | 1.47∘ | 1.36 |
| 3 | 1.56 | 1.61 | 1.64 | 1.69 | 1.65 | **1.49** | 1.63 | 1.68 | 1.59 | 1.67 | **1.52** |
| 4 | 1.42 | 1.79∘ | 1.78• | 1.85• | 1.77• | 1.49 | 1.69∘ | 1.87• | 1.71· | 1.85• | 1.64∘ |
| 5 | 1.41 | 1.83• | 1.79• | 1.88• | 1.79• | 1.41 | 1.76∘ | 1.95• | 1.73∘ | 2.00• | 1.63∘ |
| | | | | | jump off 1 | | | | | | |
| 2 | 1.27 | 1.38· | 1.42 | 1.42 | 1.41 | 1.32 | 1.46 | 1.41∘ | 1.36 | 1.38 | 1.28 |
| 3 | 1.56 | **1.55** | 1.59 | 1.60 | 1.57 | **1.47** | 1.59 | 1.63 | **1.47** | 1.65 | **1.48** |
| 4 | 1.42 | 1.52 | 1.61∘ | 1.63∘ | 1.59· | 1.42 | 1.53 | 1.71∘ | 1.51 | 1.65· | 1.50 |
| 5 | 1.41 | 1.60∘ | 1.71• | 1.74• | 1.68• | 1.44 | 1.60· | 1.81• | 1.62∘ | 1.88• | 1.56∘ |
| | | | | | jump off 2 | | | | | | |
| 3 | 1.56 | **1.50** | **1.49** | **1.49** | **1.46** | **1.40** | **1.43** | **1.54** | **1.47** | **1.53** | **1.48** |
| 4 | 1.42 | 1.44 | 1.56· | 1.58· | 1.54 | **1.40** | 1.52 | 1.60∘ | 1.42 | 1.59∘ | 1.45 |
| 5 | 1.41 | 1.44 | 1.60• | 1.64• | 1.59∘ | **1.38** | 1.47 | 1.66• | 1.47 | 1.65∘ | 1.47 |
| | | | | | jump off 3 | | | | | | |
| 4 | 1.42 | 1.49 | 1.50∘ | 1.50∘ | 1.48· | **1.41** | 1.47 | 1.53∘ | 1.45 | 1.49 | 1.45 |
| 5 | 1.41 | 1.50 | 1.55∘ | 1.56∘ | 1.51· | **1.37** | 1.47 | 1.58• | 1.45 | 1.54 | 1.46 |

Notes: Each column, 2–11, reports the root mean square prediction error (RMSPE) for forecast labelled at top for the current quarter (hor= 0) through 5 quarters into the future. For forecasts other than Greenbook, the data on which the forecast is based is the real-time data augmented to bring the data up to the jumping off point using the Greenbook forecast. Bold indicates that the alternative forecast has smaller RMSPE than Greenbook; •,∘,· indicate that the difference between the RMSPE for the model and Greenbook is statistically significant at the 1, 5, or 10 percent level, respectively, based on the bootstrap $p$-values for the DM test as described in the text.

Table 5c: RMSPE of output growth forecasts, 1984-2000

| hor. | GB | RW | IAR | DAR | EWA | BMA | FAAR | FAV | IFAV | DF | UCSV |
|---|---|---|---|---|---|---|---|---|---|---|---|
| | | | | | jump off -1 | | | | | | |
| 0 | 1.75 | 2.26· | 1.91 | 1.91 | 1.91 | 1.93 | 2.27· | 2.20· | 2.20 | 2.30 | 2.03 |
| 1 | 2.12 | 2.20 | **1.85** | **1.82** | **1.86** | **1.97** | 2.40 | 2.25 | 2.29 | 2.44 | **1.89** |
| 2 | 2.01 | 2.49· | **1.96** | **1.94** | **1.90** | **1.98** | 2.20 | 2.06 | 2.46 | 2.15 | 2.20 |
| 3 | 2.15 | 2.54 | **2.09** | **2.04** | **2.03** | 2.20 | 2.35 | 2.27 | 2.60 | 2.18 | 2.31 |
| 4 | 2.08 | 2.75∘ | **2.05** | **2.07** | **2.05** | 2.18 | 2.24 | 2.13 | 2.56 | 2.20 | 2.33 |
| 5 | 2.08 | 2.97● | 2.12 | **2.03** | **2.01** | 2.08 | 2.08 | 2.08 | 2.73· | 2.14 | 2.53 |
| | | | | | jump off 0 | | | | | | |
| 1 | 2.12 | **2.04** | **1.78** | **1.78** | **1.79** | **1.80** | **2.04** | **1.95** | 2.12 | **2.11** | **1.79∘** |
| 2 | 2.01 | 2.23 | **1.88** | **1.86** | **1.83** | **1.80** | 2.09 | **1.96** | 2.20 | 2.19 | 2.07 |
| 3 | 2.15 | 2.36 | **2.07** | **2.05** | **2.04** | **2.12** | 2.36 | 2.23 | 2.49 | 2.24 | 2.19 |
| 4 | 2.08 | 2.57· | **2.05** | **2.02** | **2.00** | 2.12 | 2.18 | 2.14 | 2.59 | 2.23 | 2.28 |
| 5 | 2.08 | 2.71● | 2.12 | **2.06** | **2.03** | 2.17 | 2.14 | 2.12 | 2.59 | 2.26 | 2.44 |
| | | | | | jump off 1 | | | | | | |
| 2 | 2.01 | 2.31 | **1.88** | **1.88** | **1.83** | **1.81** | **1.89** | **1.85** | **1.93** | **1.99** | **1.91** |
| 3 | 2.15 | 2.48 | **2.07** | **2.06** | **2.06** | **2.04** | 2.29 | 2.17 | 2.44 | 2.25 | **2.10** |
| 4 | 2.08 | 2.44∘ | **2.05** | **2.04** | **2.03** | **2.06** | 2.26 | 2.15 | 2.40 | 2.17 | 2.13 |
| 5 | 2.08 | 2.54∘ | 2.12 | 2.11 | 2.09 | 2.24 | 2.22 | 2.17 | 2.48 | 2.35 | 2.25 |
| | | | | | jump off 2 | | | | | | |
| 3 | 2.15 | 2.17 | **2.07** | **2.07** | **2.04** | **1.97** | **2.07** | **1.99** | 2.31 | **2.09** | **2.01** |
| 4 | 2.08 | 2.16 | **2.03** | **2.03** | **2.01** | **1.96** | 2.19 | 2.10 | 2.32 | 2.12 | **2.04** |
| 5 | 2.08 | 2.21 | 2.11 | 2.12 | 2.11 | 2.15 | 2.30 | 2.18 | 2.45 | 2.27 | 2.12 |
| | | | | | jump off 3 | | | | | | |
| 4 | 2.08 | 2.17 | **2.03** | **2.03** | **1.98** | **1.93** | **1.98** | **1.94** | 2.18 | **2.02** | **2.00** |
| 5 | 2.08 | 2.17 | 2.08 | 2.10 | **2.07** | **1.99** | 2.20 | 2.10 | 2.31 | 2.22 | **2.06** |

Notes: Each column, 2–11, reports the root mean square prediction error (RMSPE) for forecast labelled at top for the current quarter (hor= 0) through 5 quarters into the future. For forecasts other than Greenbook, the data on which the forecast is based is the real-time data augmented to bring the data up to the jumping off point using the Greenbook forecast. Bold indicates that the alternative forecast has smaller RMSPE than Greenbook; ●,∘,· indicate that the difference between the RMSPE for the model and Greenbook is statistically significant at the 1, 5, or 10 percent level, respectively, based on the bootstrap $p$-values for the DM test as described in the text.

Table 6a: Bias and standard deviation of deflator inflation forecasts

| hor. | GB | RW | IAR | DAR | EWA | BMA | FAAR | FAV | IFAV | DF | UCSV |
|------|------|-------|------|------|------|-------|------|------|-------|------|-------|
| | | | | | bias, 1979–2000 | | | | | | |
| 0 | 0.07 | -0.20 | 0.12 | 0.12 | 0.11 | -0.02 | 0.26 | 0.49 | -0.18 | 0.40 | -0.06 |
| 1 | 0.23 | -0.06 | 0.39 | 0.38 | 0.37 | 0.06 | 0.45 | 0.86 | -0.11 | 0.65 | 0.07 |
| 2 | 0.27 | 0.00 | 0.45 | 0.42 | 0.39 | 0.07 | 0.35 | 1.10 | -0.05 | 0.77 | 0.13 |
| 3 | 0.35 | 0.12 | 0.69 | 0.71 | 0.67 | 0.23 | 0.53 | 1.38 | 0.07 | 1.07 | 0.25 |
| 4 | 0.37 | 0.18 | 0.84 | 0.86 | 0.81 | 0.31 | 0.47 | 1.58 | 0.16 | 1.33 | 0.31 |
| 5 | 0.53 | 0.33 | 1.05 | 1.12 | 1.04 | 0.47 | 0.68 | 1.86 | 0.33 | 1.49 | 0.46 |
| | | | | | std. dev., 1979–2000 | | | | | | |
| 0 | 0.81 | 1.30 | 1.08 | 1.08 | 1.06 | 1.06 | 1.16 | 1.19 | 1.23 | 1.24 | 1.13 |
| 1 | 0.81 | 1.50 | 1.20 | 1.20 | 1.17 | 1.17 | 1.15 | 1.33 | 1.42 | 1.50 | 1.24 |
| 2 | 0.87 | 1.27 | 1.08 | 1.09 | 1.06 | 1.11 | 1.17 | 1.23 | 1.18 | 1.42 | 1.18 |
| 3 | 0.97 | 1.29 | 1.13 | 1.15 | 1.13 | 1.08 | 1.39 | 1.37 | 1.18 | 1.42 | 1.27 |
| 4 | 0.99 | 1.39 | 1.18 | 1.20 | 1.23 | 1.16 | 1.59 | 1.55 | 1.42 | 1.36 | 1.36 |
| 5 | 1.19 | 1.68 | 1.37 | 1.38 | 1.40 | 1.31 | 1.72 | 1.57 | 1.62 | 1.55 | 1.60 |
| | | | | | bias, 1984–2000 | | | | | | |
| 0 | 0.09 | -0.15 | 0.15 | 0.15 | 0.13 | 0.10 | 0.13 | 0.38 | -0.13 | 0.58 | -0.09 |
| 1 | 0.20 | -0.04 | 0.41 | 0.41 | 0.37 | 0.24 | 0.31 | 0.74 | -0.02 | 0.94 | 0.02 |
| 2 | 0.24 | -0.05 | 0.46 | 0.44 | 0.39 | 0.23 | 0.23 | 0.91 | 0.00 | 1.06 | 0.01 |
| 3 | 0.29 | -0.02 | 0.61 | 0.64 | 0.58 | 0.29 | 0.34 | 1.12 | 0.07 | 1.33 | 0.04 |
| 4 | 0.32 | 0.00 | 0.75 | 0.77 | 0.71 | 0.31 | 0.36 | 1.29 | 0.14 | 1.61 | 0.06 |
| 5 | 0.42 | 0.08 | 0.92 | 1.00 | 0.91 | 0.37 | 0.49 | 1.52 | 0.28 | 1.73 | 0.14 |
| | | | | | std. dev., 1984–2000 | | | | | | |
| 0 | 0.69 | 1.12 | 0.95 | 0.95 | 0.94 | 0.93 | 1.00 | 1.05 | 1.07 | 0.96 | 0.98 |
| 1 | 0.76 | 1.32 | 1.05 | 1.05 | 1.02 | 0.94 | 1.00 | 1.13 | 1.20 | 1.13 | 1.09 |
| 2 | 0.77 | 1.23 | 0.96 | 0.96 | 0.94 | 0.91 | 0.94 | 1.02 | 1.05 | 1.14 | 1.01 |
| 3 | 0.89 | 1.08 | 0.98 | 0.97 | 0.95 | 0.92 | 1.11 | 1.14 | 0.97 | 1.09 | 1.04 |
| 4 | 0.84 | 1.16 | 1.02 | 1.02 | 0.99 | 0.91 | 1.09 | 1.04 | 1.06 | 0.97 | 1.07 |
| 5 | 1.06 | 1.46 | 1.19 | 1.17 | 1.15 | 1.07 | 1.30 | 1.20 | 1.28 | 1.16 | 1.28 |

Notes: Each column, 2–11, reports bias or standard deviation of the forecast for the model labelled at the top. All entries are for jumping off point -1. See also notes to Table 3.

Table 6b: Bias and standard deviation of CPI inflation forecasts

| hor. | GB | RW | IAR | DAR | EWA | BMA | FAAR | FAV | IFAV | DF | UCSV |
|---|---|---|---|---|---|---|---|---|---|---|---|
| | | | | | bias, 1979–2000 | | | | | | |
| 0 | 0.13 | 0.04 | 0.31 | 0.31 | 0.25 | -0.10 | -0.02 | 0.18 | 0.05 | 0.14 | 0.18 |
| 1 | 0.27 | 0.07 | 0.49 | 0.48 | 0.40 | -0.16 | 0.14 | 0.42 | 0.14 | 0.28 | 0.20 |
| 2 | 0.38 | 0.23 | 0.62 | 0.58 | 0.51 | -0.14 | 0.20 | 0.83 | 0.42 | 0.62 | 0.36 |
| 3 | 0.36 | 0.40 | 0.97 | 0.93 | 0.85 | 0.06 | 0.35 | 1.24 | 0.71 | 0.87 | 0.53 |
| 4 | 0.37 | 0.39 | 1.12 | 1.08 | 0.98 | 0.10 | 0.64 | 1.42 | 0.82 | 0.93 | 0.52 |
| 5 | 0.54 | 0.56 | 1.31 | 1.31 | 1.21 | 0.58 | 1.02 | 1.75 | 1.09 | 1.17 | 0.69 |
| | | | | | std. dev., 1979–2000 | | | | | | |
| 0 | 0.81 | 2.06 | 1.73 | 1.73 | 1.71 | 1.62 | 1.66 | 1.85 | 1.94 | 1.76 | 1.92 |
| 1 | 1.64 | 2.21 | 1.74 | 1.76 | 1.74 | 1.56 | 1.89 | 2.12 | 2.07 | 1.92 | 1.91 |
| 2 | 1.75 | 1.95 | 1.75 | 1.87 | 1.85 | 1.58 | 1.95 | 1.94 | 1.69 | 1.84 | 1.90 |
| 3 | 1.80 | 2.49 | 2.14 | 2.18 | 2.15 | 1.87 | 2.26 | 2.31 | 2.15 | 1.94 | 2.37 |
| 4 | 1.56 | 2.55 | 1.94 | 1.95 | 1.96 | 1.85 | 2.42 | 2.38 | 2.38 | 2.10 | 2.31 |
| 5 | 1.55 | 2.39 | 1.84 | 1.78 | 1.75 | 2.14 | 2.21 | 2.06 | 2.09 | 1.85 | 2.28 |
| | | | | | bias, 1984–2000 | | | | | | |
| 0 | 0.04 | -0.01 | 0.24 | 0.24 | 0.18 | -0.07 | -0.08 | 0.06 | 0.05 | 0.21 | 0.04 |
| 1 | 0.10 | -0.02 | 0.45 | 0.46 | 0.38 | -0.05 | 0.08 | 0.23 | 0.16 | 0.53 | 0.03 |
| 2 | 0.19 | 0.00 | 0.50 | 0.54 | 0.47 | -0.07 | 0.18 | 0.47 | 0.34 | 0.81 | 0.05 |
| 3 | 0.16 | 0.03 | 0.69 | 0.76 | 0.66 | -0.02 | 0.28 | 0.73 | 0.54 | 0.96 | 0.08 |
| 4 | 0.32 | 0.09 | 0.93 | 1.04 | 0.93 | 0.14 | 0.56 | 0.99 | 0.75 | 1.19 | 0.14 |
| 5 | 0.38 | 0.14 | 1.04 | 1.19 | 1.07 | 0.21 | 0.71 | 1.21 | 0.93 | 1.32 | 0.19 |
| | | | | | std. dev., 1984–2000 | | | | | | |
| 0 | 0.49 | 1.62 | 1.45 | 1.45 | 1.45 | 1.36 | 1.47 | 1.53 | 1.48 | 1.44 | 1.44 |
| 1 | 1.15 | 1.82 | 1.42 | 1.44 | 1.42 | 1.35 | 1.43 | 1.71 | 1.57 | 1.73 | 1.57 |
| 2 | 1.26 | 1.71 | 1.54 | 1.61 | 1.59 | 1.41 | 1.58 | 1.58 | 1.53 | 1.47 | 1.57 |
| 3 | 1.56 | 2.08 | 1.86 | 1.95 | 1.91 | 1.69 | 1.90 | 1.82 | 1.86 | 1.65 | 1.87 |
| 4 | 1.39 | 2.06 | 1.61 | 1.71 | 1.67 | 1.48 | 1.79 | 1.77 | 1.71 | 1.65 | 1.81 |
| 5 | 1.37 | 1.93 | 1.56 | 1.59 | 1.59 | 1.44 | 1.77 | 1.57 | 1.69 | 1.62 | 1.75 |

Notes: Each column, 2–11, reports bias or standard deviation of the forecast for the model labelled at the top. All entries are for jumping off point -1. See also notes to Table 3.

Table 6c: Bias and standard deviation of output growth forecasts

| hor. | GB | RW | IAR | DAR | EWA | BMA | FAAR | FAV | IFAV | DF | UCSV |
|------|------|------|------|------|------|------|------|------|------|------|------|
| | | | | | bias, 1979–2000 | | | | | | |
| 0 | -0.41 | -0.29 | 0.13 | 0.13 | 0.21 | 0.02 | 0.80 | 0.62 | 0.85 | 0.13 | -0.10 |
| 1 | -0.34 | -0.04 | 0.46 | 0.44 | 0.55 | 0.43 | 1.36 | 1.05 | 1.24 | 0.67 | 0.16 |
| 2 | -0.34 | -0.23 | 0.36 | 0.34 | 0.43 | 0.56 | 1.13 | 0.89 | 1.35 | 0.24 | -0.03 |
| 3 | -0.19 | -0.01 | 0.62 | 0.48 | 0.55 | 0.86 | 1.20 | 1.02 | 1.59 | 0.33 | 0.19 |
| 4 | -0.21 | -0.07 | 0.58 | 0.56 | 0.62 | 0.91 | 0.94 | 0.82 | 1.54 | 0.33 | 0.13 |
| 5 | 0.04 | 0.07 | 0.73 | 0.68 | 0.74 | 1.02 | 1.07 | 0.83 | 1.62 | 0.57 | 0.27 |
| | | | | | std. dev., 1979–2000 | | | | | | |
| 0 | 2.14 | 3.30 | 2.78 | 2.78 | 2.71 | 2.74 | 2.64 | 2.67 | 2.61 | 2.76 | 2.75 |
| 1 | 2.74 | 3.58 | 2.72 | 2.73 | 2.59 | 2.46 | 2.46 | 2.42 | 2.53 | 2.59 | 2.72 |
| 2 | 2.71 | 3.82 | 2.81 | 2.89 | 2.78 | 2.81 | 2.65 | 2.55 | 2.65 | 2.92 | 2.97 |
| 3 | 2.77 | 3.41 | 2.54 | 2.55 | 2.49 | 2.53 | 2.57 | 2.56 | 2.66 | 2.62 | 2.84 |
| 4 | 2.57 | 4.09 | 2.63 | 2.66 | 2.56 | 2.44 | 2.39 | 2.48 | 2.71 | 2.66 | 2.97 |
| 5 | 2.42 | 3.63 | 2.57 | 2.66 | 2.61 | 2.68 | 2.36 | 2.47 | 3.25 | 2.71 | 2.91 |
| | | | | | bias, 1984–2000 | | | | | | |
| 0 | -0.28 | 0.01 | 0.13 | 0.13 | 0.27 | 0.29 | 1.04 | 0.90 | 0.89 | 0.16 | -0.13 |
| 1 | -0.19 | 0.26 | 0.38 | 0.33 | 0.52 | 0.74 | 1.43 | 1.26 | 1.23 | 0.69 | 0.12 |
| 2 | -0.14 | 0.23 | 0.39 | 0.32 | 0.49 | 0.74 | 1.37 | 1.22 | 1.39 | 0.29 | 0.09 |
| 3 | -0.04 | 0.47 | 0.63 | 0.50 | 0.65 | 1.14 | 1.49 | 1.35 | 1.58 | 0.34 | 0.32 |
| 4 | -0.17 | 0.36 | 0.53 | 0.40 | 0.52 | 1.04 | 1.15 | 1.10 | 1.42 | 0.33 | 0.22 |
| 5 | -0.09 | 0.39 | 0.57 | 0.39 | 0.47 | 0.77 | 1.00 | 0.99 | 1.35 | 0.35 | 0.25 |
| | | | | | std. dev., 1984–2000 | | | | | | |
| 0 | 1.73 | 2.27 | 1.91 | 1.91 | 1.90 | 1.91 | 2.02 | 2.02 | 2.03 | 2.30 | 2.03 |
| 1 | 2.12 | 2.19 | 1.81 | 1.79 | 1.79 | 1.83 | 1.93 | 1.87 | 1.94 | 2.35 | 1.89 |
| 2 | 2.01 | 2.49 | 1.93 | 1.92 | 1.85 | 1.84 | 1.73 | 1.66 | 2.04 | 2.14 | 2.21 |
| 3 | 2.16 | 2.51 | 2.00 | 1.99 | 1.93 | 1.88 | 1.83 | 1.83 | 2.08 | 2.17 | 2.29 |
| 4 | 2.08 | 2.74 | 1.99 | 2.04 | 1.99 | 1.92 | 1.93 | 1.83 | 2.13 | 2.19 | 2.33 |
| 5 | 2.09 | 2.95 | 2.05 | 2.00 | 1.96 | 1.94 | 1.84 | 1.84 | 2.38 | 2.12 | 2.53 |

Notes: Each column, 2–11, reports bias or standard deviation of the forecast for the model labelled at the top. All entries are for jumping off point -1. See also notes to Table 3.

Table 7: RMSPE of forecasts, most recent vintage

| hor. | GB | RW | IAR | DAR | EWA | BMA | FAAR | FAV | IFAV | DF | UCSV |
|------|------|------|------|------|------|------|------|------|------|------|------|
| | | | | | deflator inflation | | | | | | |
| 0 | 0.66 | 0.91 | 0.87 | 0.87 | 0.86 | 0.82 | 1.00 | 1.07 | 0.90 | 1.05 | 0.87 |
| 1 | 0.71 | 1.11 | 1.06 | 1.04 | 1.04 | 0.92 | 1.50 | 1.55 | 1.09 | 1.41 | 1.03 |
| 2 | 0.80 | 1.11 | 1.16 | 1.15 | 1.15 | 0.93 | 1.65 | 1.81 | 1.04 | 1.63 | 1.12 |
| 3 | 0.89 | 1.28 | 1.34 | 1.35 | 1.36 | 1.02 | 1.80 | 2.14 | 1.18 | 1.80 | 1.32 |
| 4 | 0.97 | 1.50 | 1.52 | 1.58 | 1.61 | 1.09 | 2.06 | 2.57 | 1.39 | 1.94 | 1.48 |
| 5 | 1.11 | 1.57 | 1.63 | 1.70 | 1.72 | 1.06 | 2.09 | 2.86 | 1.41 | 1.97 | 1.57 |
| | | | | | CPI inflation | | | | | | |
| 0 | 0.77 | 1.89 | 1.65 | 1.65 | 1.62 | 1.53 | 1.68 | 1.90 | 1.82 | 1.62 | 1.79 |
| 1 | 1.57 | 2.28 | 1.96 | 1.97 | 1.94 | 1.72 | 2.23 | 2.45 | 2.06 | 2.01 | 2.09 |
| 2 | 1.68 | 2.06 | 1.96 | 2.01 | 1.99 | 1.65 | 2.46 | 2.58 | 1.84 | 2.04 | 2.00 |
| 3 | 1.78 | 2.40 | 2.14 | 2.20 | 2.24 | 1.87 | 3.41 | 2.96 | 2.12 | 2.26 | 2.29 |
| 4 | 1.51 | 2.60 | 2.19 | 2.20 | 2.31 | 1.99 | 4.21 | 3.41 | 2.36 | 2.48 | 2.47 |
| 5 | 1.64 | 2.58 | 2.30 | 2.28 | 2.36 | 2.30 | 4.56 | 3.58 | 2.36 | 2.38 | 2.52 |
| | | | | | output growth | | | | | | |
| 0 | 2.14 | 3.44 | 2.99 | 2.99 | 2.81 | 2.74 | 2.71 | 2.63 | 2.90 | 2.84 | 3.06 |
| 1 | 3.05 | 3.86 | 3.15 | 3.13 | 2.95 | 2.84 | 3.05 | 2.95 | 3.25 | 3.01 | 3.26 |
| 2 | 3.01 | 3.95 | 2.94 | 3.00 | 2.88 | 2.91 | 3.05 | 2.96 | 3.23 | 3.16 | 3.15 |
| 3 | 3.04 | 4.23 | 2.89 | 3.01 | 2.92 | 2.98 | 3.05 | 2.93 | 3.48 | 3.12 | 3.13 |
| 4 | 2.80 | 4.46 | 2.90 | 3.01 | 2.90 | 2.88 | 3.08 | 2.88 | 3.52 | 2.96 | 3.14 |
| 5 | 2.63 | 4.13 | 2.87 | 3.04 | 2.96 | 2.93 | 3.12 | 3.00 | 3.81 | 3.03 | 2.94 |

Notes: Each column, 2–11, reports the root mean square prediction error (RMSPE) for the forecast labelled at top for the current quarter (hor$= 0$) through 5 quarters into the future. The jumping off point is $t-1$. Other than Greenbook, the forecasts are all based on a single vintage of data: our most recent vintage. See the notes to Table 3.

Table 8: RMSPE of forecasts, Stock-Watson (2005) data

| hor. | GB | RW | IAR | DAR | EWA | BMA | FAAR | FAV | IFAV | DF | UCSV |
|------|------|------|------|------|------|------|------|------|------|------|------|
| | | | | | deflator inflation | | | | | | |
| 0 | 0.66 | 0.91 | 0.87 | 0.87 | 0.85 | 0.84 | 0.94 | 0.94 | 0.94 | 1.40 | 0.87 |
| 1 | 0.71 | 1.11 | 1.06 | 1.04 | 1.03 | 1.01 | 1.16 | 1.24 | 1.19 | 1.50 | 1.03 |
| 2 | 0.80 | 1.11 | 1.16 | 1.15 | 1.12 | 1.05 | 1.20 | 1.25 | 1.17 | 1.54 | 1.12 |
| 3 | 0.89 | 1.28 | 1.34 | 1.35 | 1.28 | 1.10 | 1.30 | 1.37 | 1.25 | 1.63 | 1.32 |
| 4 | 0.97 | 1.50 | 1.52 | 1.58 | 1.49 | 1.17 | 1.47 | 1.57 | 1.45 | 1.61 | 1.48 |
| 5 | 1.11 | 1.57 | 1.63 | 1.70 | 1.59 | 1.18 | 1.49 | 1.61 | 1.49 | 1.65 | 1.57 |
| | | | | | CPI inflation | | | | | | |
| 0 | 0.77 | 1.89 | 1.65 | 1.65 | 1.57 | 1.51 | 1.52 | 1.81 | 1.67 | 1.78 | 1.79 |
| 1 | 1.57 | 2.28 | 1.96 | 1.97 | 1.90 | 1.67 | 1.81 | 2.26 | 1.99 | 1.96 | 2.09 |
| 2 | 1.68 | 2.06 | 1.96 | 2.01 | 1.94 | 1.56 | 1.87 | 1.91 | 1.89 | 1.84 | 2.00 |
| 3 | 1.78 | 2.40 | 2.14 | 2.20 | 2.12 | 1.71 | 2.06 | 2.09 | 2.07 | 2.03 | 2.29 |
| 4 | 1.51 | 2.60 | 2.19 | 2.20 | 2.17 | 1.77 | 2.22 | 2.36 | 2.34 | 2.08 | 2.47 |
| 5 | 1.64 | 2.58 | 2.30 | 2.28 | 2.21 | 1.84 | 2.27 | 2.30 | 2.36 | 2.01 | 2.52 |
| | | | | | output growth | | | | | | |
| 0 | 2.14 | 3.44 | 2.99 | 2.99 | 2.84 | 2.76 | 2.59 | 2.46 | 2.91 | 2.73 | 3.06 |
| 1 | 3.05 | 3.86 | 3.15 | 3.13 | 2.99 | 2.90 | 2.73 | 2.60 | 2.84 | 2.95 | 3.26 |
| 2 | 3.01 | 3.95 | 2.94 | 3.00 | 2.97 | 2.95 | 3.06 | 3.00 | 3.31 | 3.16 | 3.15 |
| 3 | 3.04 | 4.23 | 2.89 | 3.01 | 2.97 | 2.96 | 3.07 | 2.90 | 3.51 | 3.11 | 3.13 |
| 4 | 2.80 | 4.46 | 2.90 | 3.01 | 2.95 | 2.84 | 2.79 | 2.83 | 3.54 | 2.99 | 3.14 |
| 5 | 2.63 | 4.13 | 2.87 | 3.04 | 3.02 | 2.97 | 2.99 | 3.07 | 3.63 | 2.95 | 2.94 |

Notes: Each column, 2–11, reports the root mean square prediction error (RMSPE) for the forecast labelled at top for the current quarter (hor= 0) through 5 quarters into the future. The jumping off point is $t-1$. Other than Greenbook, the forecasts are all based on a single vintage of data: the Stock-Watson (2005) data. See the notes to Table 3.

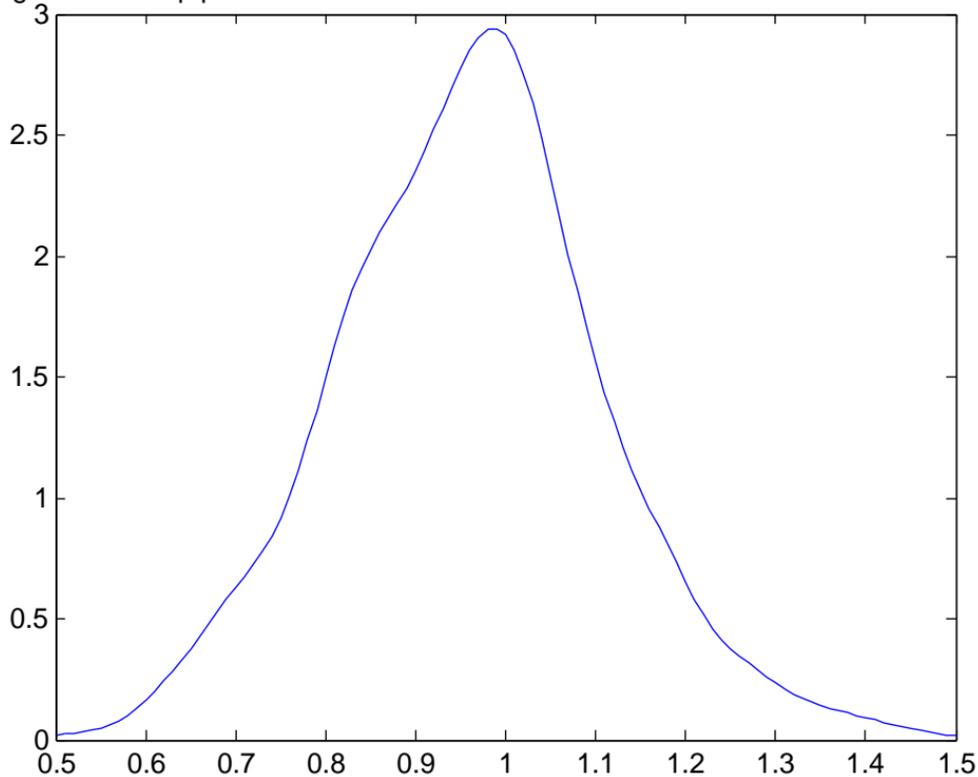Fig. 1: Bootstrap pdf for the ratio of GB to RW RMSPEs in Atkeson–Ohanian sample
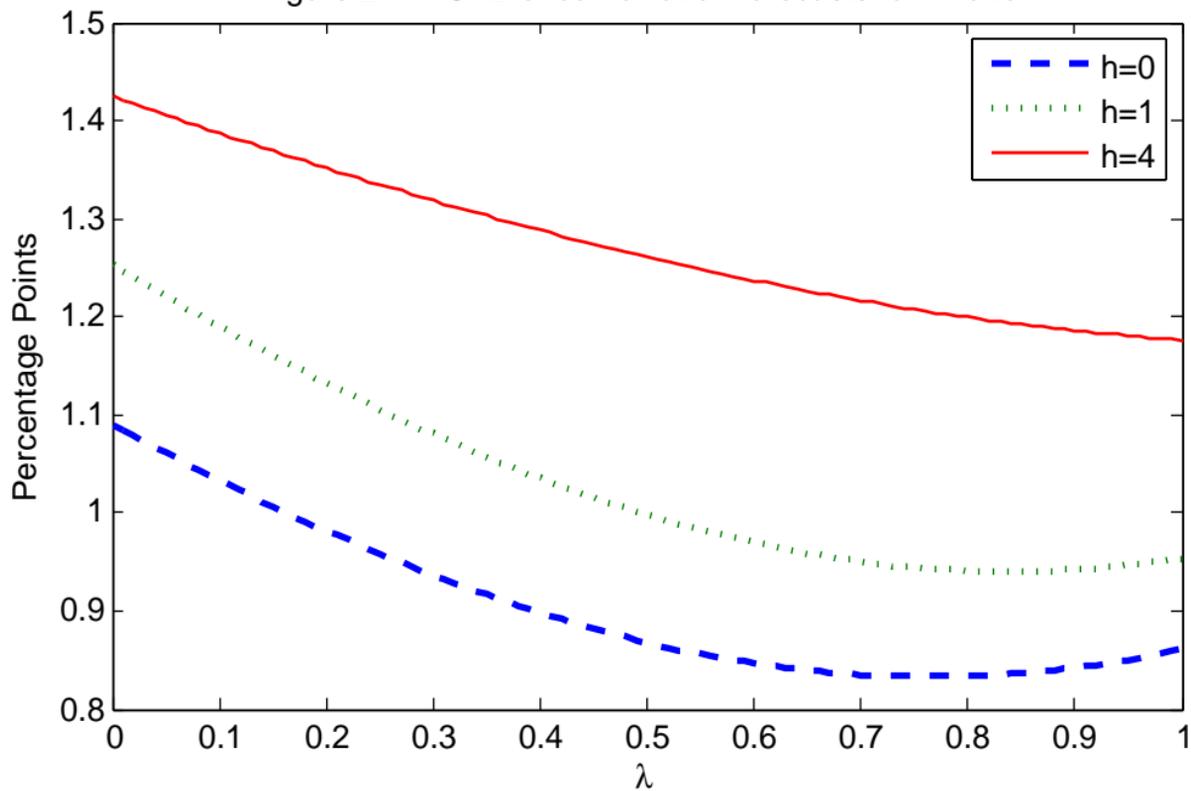
Figure 2: RMSPE of combination forecasts for inflation

Figure 3: RMSPE of combination forecasts for growth